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Bzovsky, S.; O'Hara, N.N.; Slobogean, G.P.; Sprague, S.; Axelrod, D.E.; Hoit, G.; ... ; FAITH Investigators

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# Revisiting the FAITH Trial: A Secondary Analysis Yielding Novel Insights with the Win Ratio

Sofia Bzovsky, MSc, Nathan N. O'Hara, PhD, Gerard P. Slobogean, MD, Sheila Sprague, PhD, Daniel E. Axelrod, MD, MSc, Graeme Hoit, MD, Kiara Pannozzo, MPH, Mohit Bhandari, MD, PhD, FRCSC, Marc Swionkowski, MD, Emil Schemitsch, MD, FRCSC, and Rudolf W. Poolman, MD, PhD, the FAITH Investigators

*Investigation was performed at McMaster University in Hamilton, Ontario, Canada*

**Background:** Many orthopaedic trials use any unplanned reoperation as the primary outcome, but this overlooks how patients experience those outcomes. Using a high-quality hip fracture trial, we demonstrate how the relative importance of multiple patient-important outcomes can be effectively incorporated into data analysis, providing a more comprehensive understanding of treatment impact.

**Methods:** This secondary analysis of the Fixation using Alternative Implants for the Treatment of Hip Fracture (FAITH) trial included 1,079 patients aged 50 years or older with a low-energy femoral neck fracture who were randomly assigned to treatment with a sliding hip screw or cancellous screws. The original trial used unplanned revision surgery as the primary outcome. Our primary analysis instead used a composite outcome of all-cause mortality at 4 months, ambulation status at 10 weeks (measured by the EuroQol-5 Dimension [EQ-5D] mobility dimension), and days at home within 4 months. We assessed outcomes hierarchically using the win ratio method, comparing each patient with every other patient in the alternative treatment group in a pairwise manner. We conducted sensitivity analyses at 6 and 12 months, and subgroup analyses to explore smoking status and fracture displacement as potential effect modifiers.

**Results:** Of the 1,079 participants, 741 had EQ-5D data available for the primary analysis at 4 months, yielding 137,114 pairwise comparisons. A sliding hip screw was superior to cancellous screws in 65,158 (47.5%) comparisons, inferior to cancellous screws in 63,378 (46.2%) comparisons, and tied in 8,578 (6.3%), leading to a win ratio of 1.03 (95% confidence interval [CI] 0.86-1.23), but this difference was not statistically significant ( $p = 0.76$ ). The sensitivity analysis results were similar at 6 and 12 months. In the subgroup analysis, a sliding hip screw was superior to cancellous screws in current smokers, with a win ratio of 1.65 (95% CI 1.02-2.65) at 6 months ( $p = 0.007$ ).

**Conclusion:** This analysis approach should be considered for future orthopaedic trials as it was consistent with the FAITH primary analysis findings but yielded a more nuanced interpretation of the patients' experience and offers deeper insights into intervention effectiveness. The bounds of the 95% CI for the primary outcome were within many standard definitions of equivalence, suggesting surgeons can assume similar patient-important outcomes with either treatment.

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## Introduction

Many orthopaedic trials use unplanned reoperation as the primary outcome, but this fails to delineate between minor and major reoperations, and does not convey how patients experience those outcomes. The win ratio is a statistical approach used in clinical trials to analyze hierarchical composite outcomes, providing a ranked evaluation of multiple components (e.g., survival, quality of life, and functional status)<sup>1</sup>. Unlike traditional methods such as Kaplan-Meier survival analysis, log-rank test, and Cox proportional hazards regression, the win ratio incorporates the relative importance of each outcome in a *predefined* hierarchy, comparing all possible pairs of participants across treatment groups. In contrast to conventional methods, the win ratio can analyze composites composed of time-to-event, recurrent events, continuous, and categorical outcomes<sup>1</sup>. This allows for a more comprehensive interpretation of data and often greater statistical power to identify and quantify a treatment difference by incorporating all available information within the component outcomes.<sup>2</sup>

Using data from the high-quality hip fracture trial, Fixation using Alternative Implants for the Treatment of Hip Fractures (FAITH)<sup>3</sup>, we aimed to demonstrate how the relative importance of multiple outcomes can effectively be incorporated into data analyses by separately evaluating the clinical status of these hip fracture patients, providing a more comprehensive understanding of treatment impact. Our goal is not to diminish the value of this trial but to demonstrate that alternative outcome measures and analytical approaches can yield additional information for the results and potentially enhance statistical power. By adopting this new method of analysis, we hope to offer more comprehensive insights that could guide policymakers and improve the design and interpretation of future orthopaedic trials.

## Methods

### FAITH Study Overview

Between March 3, 2008, and March 31, 2014, the FAITH trial (ClinicalTrials.gov NCT00761813) enrolled 1,079 patients from 81 clinical centers in the United States, Canada, Australia, the Netherlands, Norway, Germany, the United Kingdom, and India, who were at least 50 years old with a low-energy femoral neck fracture treated by fracture fixation with a sliding hip screw or cancellous screws. The primary outcome was unplanned revision surgery to promote healing, relieve pain, treat infection, or improve function within 24 months of fracture fixation. Ethics approval for this secondary analysis was obtained from the Hamilton Integrated Research Ethics Board (#18542).

### Outcomes

For this secondary analysis of the FAITH trial, we constructed a ranked composite outcome that included (1) all-cause mortality, (2) ambulation status at one follow-up time point assessed using the EuroQol-5 Dimension (EQ-5D) mobility dimension<sup>4</sup>, and (3) number of days at home (see Appendix Figure S1). Higher priority was given to components that are

clinically more important, with mortality considered the most critical outcome, followed by ambulation status, and then days at home. The ranking of these components was supported by patient preference research in this study population<sup>5-8</sup>.

The FAITH trial collected the length of stay for hospital admissions for the initial surgery, reoperations, and adverse events throughout the 24-month follow-up period. Using these data points, the days in hospital were calculated for each participant and subtracted from the days alive within the 24-month follow-up period to calculate the number of days at home. We present a breakdown of the average length of hospital stay in FAITH trial participants by reoperation type and reason for reoperation in the Appendix. In the FAITH trial, reoperations that were classified as study events included implant removal; implant exchange—total hip arthroplasty, hemiarthroplasty, or internal fixation; soft tissue procedure; and any other event as determined by the adjudication committee (proximal femoral osteotomy). The reasons for reoperations included painful hardware, implant failure, avascular necrosis, non-union, deep infection, superficial infection, hip instability, intractable pain due to wear of the acetabulum, periprosthetic femur fracture, or hip dislocation.

### Statistical Analysis

Baseline demographics and fracture characteristics were analyzed using descriptive statistics reported as count and percentage or mean and standard deviation (SD) or median and interquartile range, depending on the data distribution.

The primary analyses followed the intention-to-treat principle, analyzing participants in the group to which they were randomly assigned. We hierarchically assessed mortality at 4 months, followed by ambulation status (EQ-5D mobility dimension) at 10-week postrandomization, and the number of days at home within 4-month postrandomization using the win ratio method. The win ratio method is based on the principle that each patient in a clinical trial is compared with every other patient assigned to the alternative treatment in a pairwise manner. Higher importance is given to higher-ranked components of the composite outcome. The pairwise comparison proceeds in a *predefined* hierarchical fashion, starting with all-cause mortality, followed by ambulation status (EQ-5D mobility dimension), and then days at home when patients cannot be differentiated based on a higher ranked comparison. The all-cause mortality component was included as a time-to-event analysis, assuming an earlier event was worse than a later occurrence. Participants with unknown mortality status at the time of assessment were censored at their last known observation. If a participant died within an assessment window, rendering their ambulation status unknown, they were assigned the lowest ambulation level (confined to bed) to not be dropped from the analysis. Finally, the treatment groups were assigned a win, loss, or tie in each pairwise comparison. Initially, the pairs were compared for time until death, truncated at 4 months. If both participants died, the “winner” of the comparison was the one who had a longer time between the time of

randomization and the date of death. If the match was tied (both participants died within the same follow-up time or both remained alive until the 4-month visit), the pair were then compared for ambulation status. Finally, if a second tie occurred, participants were compared for days at home, and the participant with the most days at home was declared the “winner.” The win ratio is the number of wins in one treatment group divided by the number of wins in the other treatment group with a 95% confidence interval (CI) and p-value calculated using the methods described by Bebu and Lachin<sup>9</sup>. A win ratio greater than 1 indicated a better outcome in the sliding hip screw group.

We conducted 2 sensitivity analyses that followed the primary analysis methods but changed the timing of outcomes:

1. We hierarchically assessed mortality at 6 months, followed by ambulation status (EQ-5D mobility dimension) at 6-month postrandomization, and number of days at home across 6-month postrandomization.

2. We hierarchically assessed mortality at 12 months, followed by ambulation status (EQ-5D mobility dimension) at 12-month postrandomization, and number of days at home across 12-month postrandomization.

We also conducted subgroup analyses investigating smoking status and fracture displacement as possible effects modifier at 4, 6, and 12 months. Smoking status and fracture displacement were included as subgroups in the primary FAITH study<sup>3</sup>.

For all analyses, the threshold for statistical significance was  $p < 0.05$ . We did not adjust the alpha for multiple comparisons. All analyses were conducted in R (version 4.4.2; R Foundation for Statistical Computing).

## Results

### Demographics and Fracture Characteristics

Of the 1,079 participants enrolled in the FAITH trial, 741 were included in the primary analysis. Most of these participants were female (63.7%) and White (94.3%), with a mean age of 73.2 years (SD 11.5 years). Half of the participants had a normal body mass index (18.5–24.9), and most were either former or nonsmokers (82.5%). Nearly all injuries were sustained from a fall (97.0%), and most fractures were subcapital (62.2%), undisplaced (70.3%), and of Pauwels classification Type II (62.1%) (Table I).

### Primary Analysis

Of the 1,079 participants, 741 had EQ-5D data available for the primary analysis at 4 months, yielding 137,114 pairwise comparisons (383 sliding hip screw participants  $\times$  358 cancellous screw participants). A sliding hip screw was found to be superior to cancellous screws in 65,158 (47.5%) comparisons, inferior to cancellous screws in 63,378 (46.2%) comparisons, and tied in 8,578 (6.3%), leading to a win ratio of 1.03 (95% CI 0.86–1.23), but this difference was not statistically significant ( $p = 0.76$ ). Mortality (5.5% vs. 7.3%), ambulation status (no problems walking: 18.3% vs. 16.2%, some problems walking: 73.6% vs. 74.9%, confined to bed: 8.1% vs. 8.9%), and days at home (111.7

**TABLE I Demographics and Fracture Characteristics**

	Sliding Hip Screw N = 383	Cancellous Screws N = 358
Age (yrs), mean (SD)	73.4 (11.6)	73.0 (11.5)
Sex, n (%)		
Male	144 (37.6)	125 (34.9)
Female	239 (62.4)	233 (65.1)
Body mass index, n (%)		
Underweight (<18.5)	26 (6.8)	24 (6.7)
Normal weight (18.5–24.9)	193 (50.4)	183 (51.1)
Overweight (25–29.9)	121 (31.6)	109 (30.4)
Obese (30–39.9)	34 (8.9)	35 (9.8)
Morbidly obese ( $\geq 40$ )	6 (1.6)	2 (0.6)
Did not disclose	3 (0.8)	5 (1.4)
Race/ethnicity, n (%)		
Indigenous	1 (0.3)	1 (0.3)
South Asian	4 (1.0)	3 (0.8)
East Asian	4 (1.0)	3 (0.8)
Black	15 (3.9)	8 (2.2)
Hispanic or Latin	1 (0.3)	1 (0.3)
White	357 (93.2)	342 (95.5)
Did not disclose	1 (0.3)	0 (0.0)
Smoking history, n (%)		
Former or nonsmoker	315 (82.2)	296 (82.7)
Current smoker	67 (17.5)	61 (17.0)
Did not disclose	1 (0.3)	1 (0.3)
Mechanism of injury, n (%)		
Fall	369 (96.3)	350 (97.8)
Spontaneous fracture	11 (2.9)	4 (1.1)
Other low-energy trauma	1 (0.3)	3 (0.8)
Did not disclose	2 (0.5)	1 (0.3)
Level of fracture line, n (%)		
Subcapital	259 (67.6)	258 (72.1)
Midcervical	109 (28.5)	88 (24.6)
Basal	15 (3.9)	12 (3.4)
Fracture displacement, n (%)		
Displaced	117 (30.5)	103 (28.8)
Undisplaced	266 (69.5)	255 (71.2)
Pauwel classification, n (%)		
Type I	40 (10.4)	31 (8.7)
Type II	297 (77.5)	279 (77.9)
Type III	46 (12.0)	48 (13.4)

vs. 111.0) were similar between the sliding hip screw and cancellous screws groups (Fig. 1, Appendix Table S1).

### Sensitivity Analyses

At 6 months, 636 participants had EQ-5D data available for the analysis, yielding 101,075 pairwise comparisons. A sliding hip screw was found to be superior to cancellous screws in 46,673

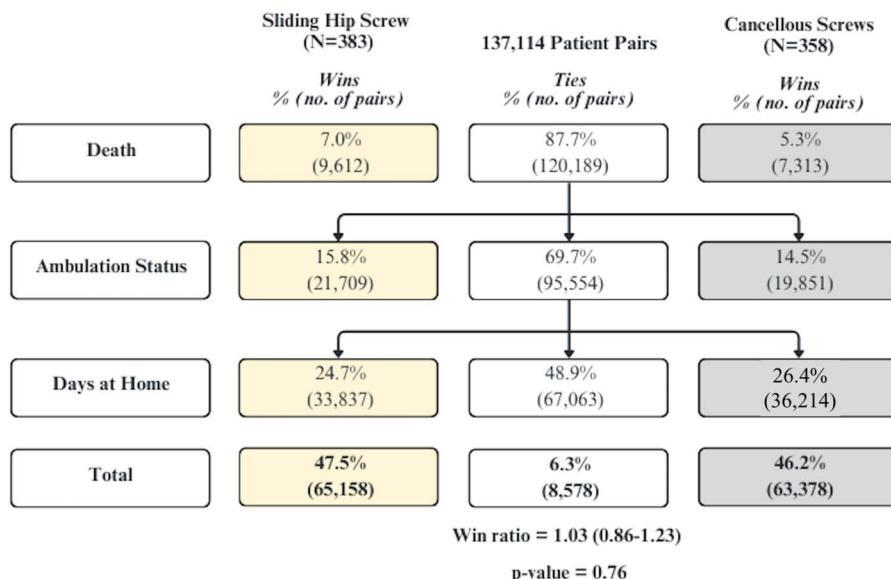


Fig. 1  
Primary analysis—win ratio results at 4 months.

(46.2%) comparisons, inferior to cancellous screws in 49,616 (49.1%) comparisons, and tied in 4,786 (4.7%), leading to a win ratio of 0.94 (95% CI 0.78-1.14), but this difference was not statistically significant ( $p = 0.53$ ) (Fig. 2, see Appendix Table S1).

At 12 months, 655 participants had EQ-5D data available for the analysis, yielding 107,226 pairwise comparisons. A sliding hip screw was found to be superior to cancellous screws in 49,751 (46.4%) comparisons, inferior to cancellous screws in 53,193 (49.6%) comparisons, and tied in 4,282 (4.0%), leading to a win ratio of 0.94 (95% CI 0.78-1.12), but this

difference also was not statistically significant ( $p = 0.48$ ) (Fig. 3, see Appendix Table S1).

#### Subgroup Analyses—Smoking Status

At 4 months, the win ratio between the sliding hip screw and cancellous screws groups in current smokers was 1.41 (95% CI 0.92-2.17), with the sliding hip screw group experiencing more wins in all outcome components compared with the cancellous screws group (overall: 55.2% vs. 39.1%), but this difference only neared statistical significance ( $p = 0.09$ ) (Fig. 4, see Appendix Table S1).

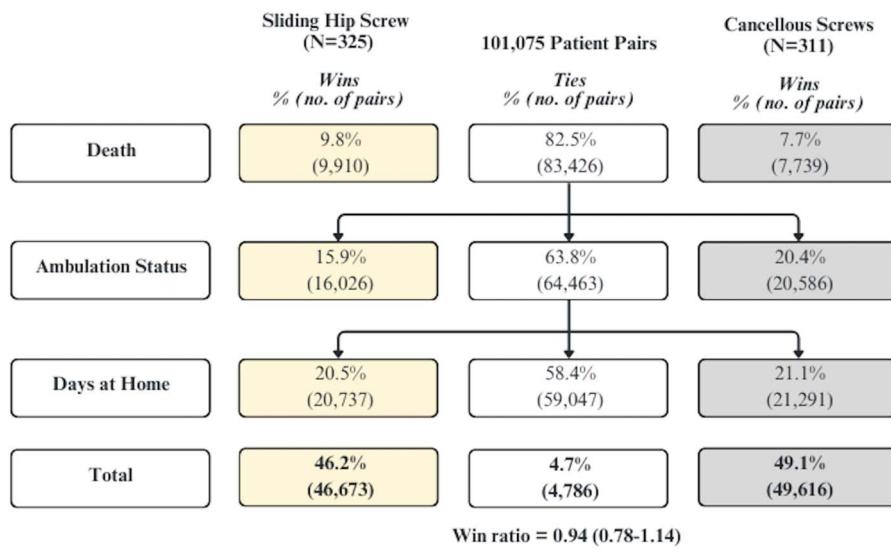


Fig. 2  
Sensitivity analysis—win ratio results at 6 months.

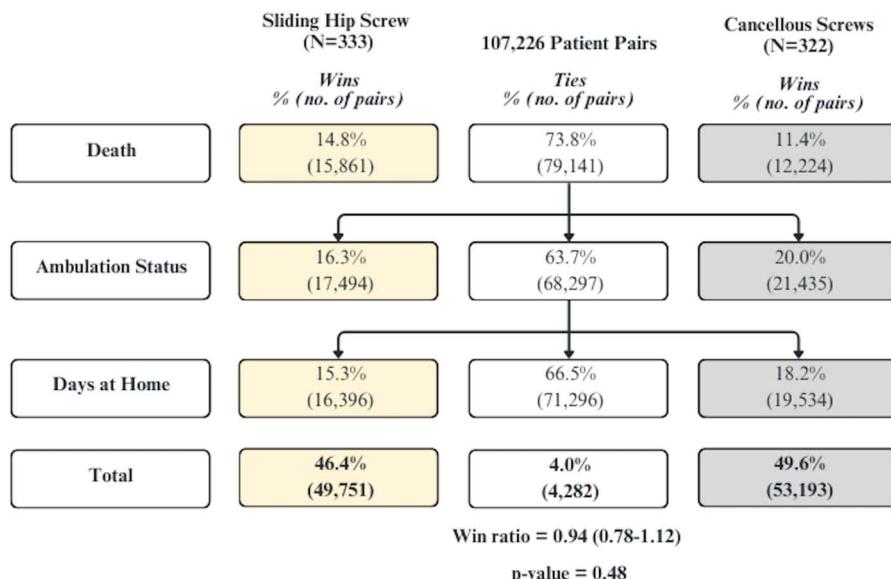


Fig. 3  
Sensitivity analysis—win ratio results at 12 months.

At 6 months, treatment with a sliding hip screw was preferred in current smokers, with a win ratio of 1.65 (95% 1.02-2.65;  $p = 0.007$ ). The sliding hip screw group won on the mortality (15.3% wins vs. 10.8%), ambulation status (25.7% wins vs. 9.1%), and days at home (18.9% wins vs. 16.5%) outcomes (Fig. 4, see Appendix Table S1).

At 12 months, the win ratio between the sliding hip screw and cancellous screws groups in current smokers was 1.29 (95% CI 0.83-2.00), with the sliding hip screw group experiencing more wins overall compared with the cancellous screws group (53.8% vs. 41.7%), but this difference did not reach statistical significance ( $p = 0.10$ ) (Fig. 4, see Appendix Table S1).

#### Subgroup Analyses—Fracture Displacement

At 4 months, the win ratio between the sliding hip screw and cancellous screws groups in those with a undisplaced fracture was 1.06 (95% CI 0.86-1.32), with the sliding hip screw group experiencing more wins overall compared with the cancellous screws group (48.4% vs. 45.4%), but this difference did not reach statistical significance ( $p = 0.53$ ) (see Appendix Table S1).

At 6 and 12 months, the win ratios between the sliding hip screw and cancellous screws groups in those with a undisplaced fracture were 0.90 (95% CI 0.72-1.13) and 0.96 (95% CI 0.77-1.20), respectively, with the sliding hip screw group experiencing more losses overall compared with the cancellous screws group (6 months: 45.2% vs. 50.3%; 12 months: 47.0% vs. 49.0%). These differences did not reach statistical significance (6 months:  $p = 0.37$ ; 12 months:  $p = 0.63$ ) (see Appendix Table S1).

#### Discussion

In this reanalysis of the FAITH trial, the results were consistent with the FAITH primary analysis findings but yielded a

more nuanced interpretation of composite data when using a hierarchical composite outcome consisting of all-cause mortality, ambulation status, and number of days at home. The bounds of the 95% CI meet many standard definitions of equivalence (between 0.8-1.25 on a relative scale according to the U.S. Food and Drug Administration guidance)<sup>10</sup>, suggesting surgeons can assume similar patient-important outcomes with either treatment.

The win ratio method has been widely adopted in cardiovascular trials over the past decade for analyzing a composite clinical hierarchy of outcomes<sup>11</sup> but has yet to be explored in orthopaedic trial research. The win ratio approach shares similarities with the conventional time-to-first event analyses (such as Cox proportional hazards models) that are typically used in orthopaedic trials in that both approaches compare outcomes between treatment groups by considering the time at which events occur. However, there are key differences in how they handle events and prioritize outcomes. Many orthopaedic trials use unplanned reoperation as the primary outcome, and if using a composite endpoint, such as reoperation and nonunion or malunion, these conventional composite endpoints do not take into account that the component events likely vary in their clinical importance and only consider the first event to occur regardless of its clinical importance<sup>1,11</sup>. Although these time-to-first event analyses provide a straightforward and statistically powerful way to compare treatment effects in clinical trials, they may not always fully capture a trial's overall conclusions<sup>12</sup>. The win ratio is useful because composite outcomes generally have a well-defined hierarchy of components that aligns with their clinical importance, allowing the outcomes to be more patient-centered. Another benefit of win statistics is its ability to also incorporate recurrent events and patient-reported outcomes (continuous or categorical) within a clinical hierarchy<sup>2,13</sup>.

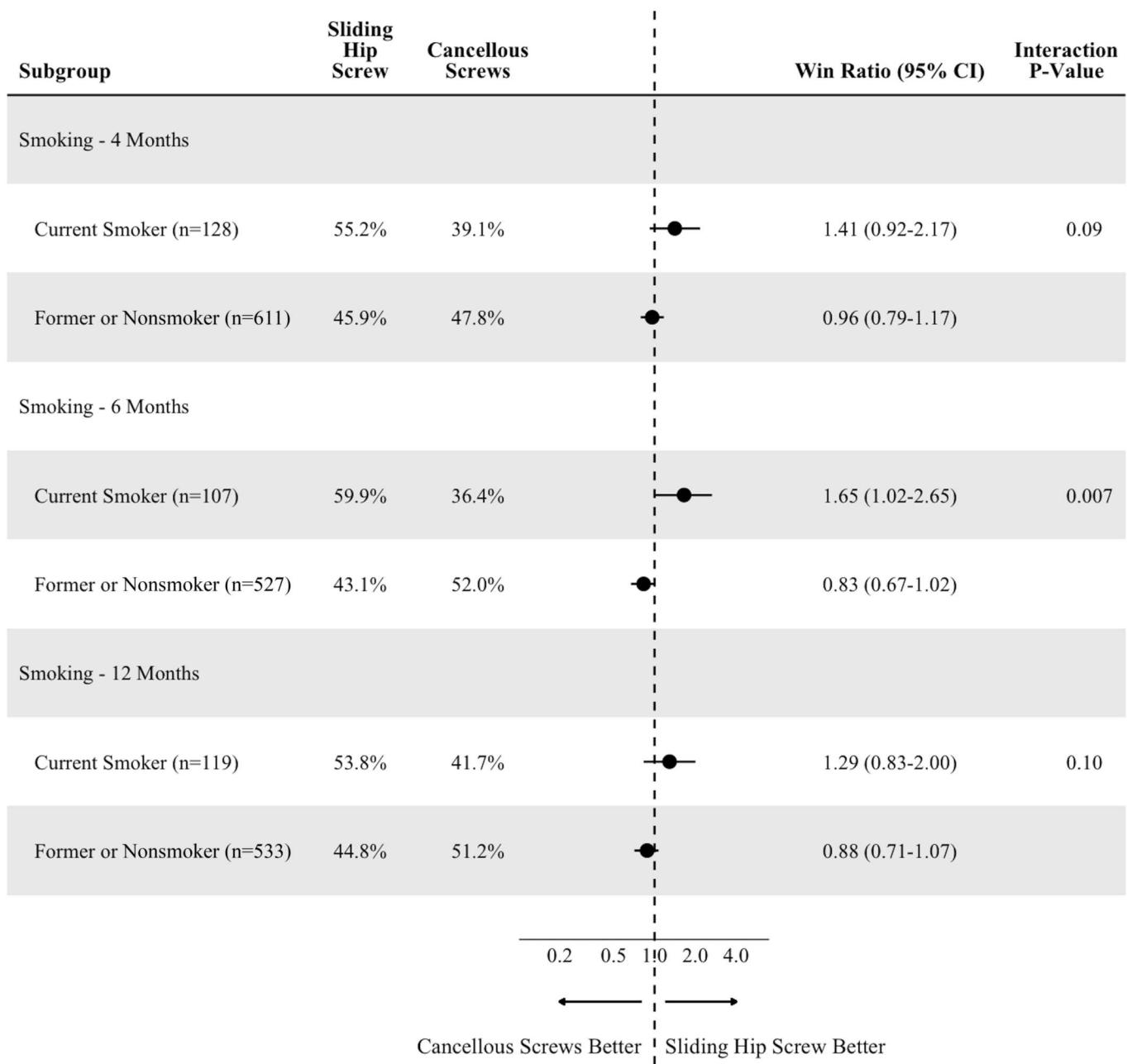


Fig. 4  
Subgroup analyses.

Overall, the win ratio approach is an appealing alternative for comparing randomized treatments due to its hierarchical structure, strong statistical power, flexibility, and ability to incorporate patient-important outcomes.

Breaking down wins and losses by component clarifies each outcome's contribution to the overall win ratio, helping determine whether the treatment effect is driven by clinically important events (e.g., mortality) or less critical outcomes (e.g., ambulation status). In our subgroup analysis, the results for current smokers at 6 months appear to be primarily influ-

enced by ambulation status, with 25.7% wins in the sliding hip screw group vs. 9.1% wins in the cancellous screw group. When examining the ambulation status breakdown, the most notable shift occurs between the “some problems walking” and “confined to bed” categories. However, a limitation of this analysis is the use of only 3 ambulation levels. Most patients fall into the “some problems walking” category, which may obscure meaningful differences. In future analyses, a more detailed 4-level classification would be beneficial: (1) ambulates without an aid, (2) ambulates with an aid, (3) ambulates with human assistance,

and (4) unable to ambulate, with ambulate defined as walking 10 feet or across a room. Adopting this refined categorization could better differentiate the “some problems walking” group, offering a more comprehensive assessment and potentially revealing differences not apparent in the current analysis.

Our analyses of the FAITH trial data demonstrate how hierarchically assessed composite endpoints are affected by the duration of follow-up. The duration of follow-up affects the number of events observed, the weight of different outcome types, and the balance of wins and losses over time<sup>1,13,14</sup>. Longer follow-up can provide a more complete picture of treatment effects, especially when an intervention has delayed benefits or if secondary outcomes (e.g., functional decline and hospitalizations) become more relevant over time. However, longer observation periods also increase the likelihood of competing risks, often exogenous to the treatments. On the other hand, shorter follow-up may limit the number of events recorded, leading to a higher proportion of ties, potentially diluting the treatment effect. For instance, in our FAITH trial subgroup analyses, the treatment effects among smokers at 4 months may not have been apparent because key complications likely driving the overall FAITH outcomes had not yet occurred within that timeframe, but became apparent at 6 months. Selecting an appropriate follow-up period is crucial to ensuring that win statistics accurately reflect treatment efficacy. Examining how the win ratio evolves over time, along with the distribution of wins and losses across different outcome tiers, helps clarify treatment effects. These patterns may vary over time between treatment groups or across different components of the outcome hierarchy, offering deeper insight into the intervention’s impact<sup>13</sup>.

This exercise demonstrates the potential of the win ratio method as a valuable alternative to conventional time-to-first event approaches in orthopaedic trials. It should be noted that although the win ratio method provides statistical verification of similarities or differences between the comparisons, the interpretation of clinical significance is left to the authors and readers. Assessing the practical or clinical importance of observed differences relies on informed judgment beyond the statistical results. Although our findings were consistent with the FAITH trial’s primary analysis, by accounting for event *predefined* hierarchies and incorporating patient-reported outcomes, the win ratio provides a patient-centered approach to assessing treatment effects. Future orthopaedic trials should consider exploring the application of win ratio statistics for evaluating treatment efficacy to enhance clinical decision-making and patient care.

## Appendix

**eA** Supporting material provided by the authors is posted with the online version of this article as a data supplement at jbjs.org (<http://links.lww.com/JBJSOA/A997>). This content was not copyedited or verified by JBJS. ■

FAITH Investigators: Writing Committee: PJ Devereaux (McMaster University), Gordon Guyatt (McMaster University), Lehana Thabane (McMaster University), Stephen D. Walter (McMaster University), Martin J. Heetveld (Spaerne Gasthuis, Haarlem), Kyle J. Jeray (Greenville Health System), Susan Liew (The Alfred), Paul Tornetta III (Boston University Medical Center), Gregory J. Della Rocca (Duke University), Richard E. Buckley (Foothills Medical Centre), Robert McCormack (Royal Columbian Hospital/Fraser Health Authority/University of British Columbia), Todd M. Oliver (Boone Hospital Center—Columbia Orthopaedic Group), Michiel J.M. Segers (St. Antonius Ziekenhuis), Amar Rangan (The James Cook University Hospital), Martin Richardson (University of Melbourne), Taryn Scott (McMaster University), Julie Agel (University of Minnesota), Alisha Garibaldi (McMaster University), Qi Zhou (McMaster University), Diane Heels-Ansdel (McMaster University), Helena Viveiros (McMaster University), Stephanie M. Zielinski (Erasmus MC, University Medical Center Rotterdam), Esther M.M. Van Lieshout (Erasmus MC, University Medical Center Rotterdam), Herman Johal (McMaster University), Birgit C. Hanusch (The James Cook University Hospital).

Steering Committee: PJ Devereaux (McMaster University), Gordon Guyatt (McMaster University), Martin J. Heetveld (Spaerne Gasthuis, Haarlem), Kyle Jeray (Greenville Health System), Susan Liew (The Alfred), Martin Richardson (University of Melbourne), Lehana Thabane (McMaster University), Paul Tornetta III (Boston University Medical Center), and Stephen D. Walter (McMaster University).

Global Methods Centre: Paula McKay (Manager); Taryn Scott, Alisha Garibaldi, Helena Viveiros, Marilyn Swinton, (Research Coordination); Mark Gichuru (Adjudication Coordination); Diane Heels-Ansdel, Qi Zhou (Statistical Analysis); Lisa Buckingham, Aravin Duraikannan (Data Management); Deborah Maddock, Nicole Simunovic (Grants Management) (McMaster University).

United States Methods Centre: Julie Agel (Research Coordination) (University of Minnesota).

Netherlands Method Centre: Martin J. Heetveld (Principal Investigator), Esther M.M. Van Lieshout (Research Coordination), Stephanie M. Zielinski (Trial Coordination) (Erasmus MC, University Medical Center Rotterdam).

United Kingdom Method Centre: Amar Rangan (Principal Investigator), Birgit C. Hanusch, Lucksy Kottam, Rachel Clarkson (Research Coordination) (The James Cook University Hospital).

Adjudication Committee: Gregory J. Della Rocca (Chair) (Duke University), Robert Haverlag (Onze Vrouwe Gasthuis), Susan Liew (The Alfred), Kyle Jeray (Greenville Health System).

Participating Clinical Sites: Canada: Robert McCormack, Kelly Apostle, Dory Boyer, Farhad Moola, Bertrand Perey, Trevor Stone, Darius Viskontas, H. Michael Lemke, Mauri Zomar, Karyn Moon, Raey Moon, Amber Oat (Royal Columbian Hospital/Fraser Health Authority/University of British Columbia); Richard E. Buckley, Paul Duffy, Robert Korley, Shannon Puloski, James Powell, Kelly Johnston, Kimberly Carcary, Melissa Lorenzen, Ross Mcckercher (Foothills Medical Centre); David Sanders, Mark MacLeod, Abdel-Rahman Lawandy, Christina Tieszer (London Health Sciences Centre); David Stephen, Hans Kreder, Richard Jenkinson, Markku Nousiainen, Terry Axelrod, John Murnaghan, Diane Nam, Veronica Wadey, Albert Yee, Katrine Milner, Monica Kunz, Wesley Ghent (Sunnybrook Health Sciences Centre); Michael D. McKee, Jeremy A. Hall, Aaron Nauth, Henry Ahn, Daniel B. Whelan, Milena R. Vicente, Lisa M. Wild, Ryan M. Khan, Jennifer T. Hidy (St. Michael’s Hospital); Chad Coles, Ross Leighton, Michael Biddulph, David Johnston, Mark Glazebrook, David Alexander, Catherine Coady, Michael Dunbar, J. David Amirault, Michael Gross, William Oxner, Gerald Reardon, C. Glen Richardson, J. Andrew Trenholm, Ivan Wong, Kelly Trask, Shelley MacDonald, Gwendolyn Dobbin (Queen Elizabeth II Health Sciences Centre); Ryan Bicknell, Jeff Yach, Davide Bardana, Gavin Wood, Mark Harrison, David Yen, Sue Lambert, Fiona Howells, Angela Ward (Human Mobility Research Centre, Queen’s University and Kingston General Hospital); Paul Zalzal, Heather Brénier, V. Naumetz, Brad Weening, Nicole Simunovic (Oakville Trafalgar Memorial Hospital); Eugene K. Wai, Steve Papp, Wade T. Gotoff, Allen Liew, Stephen P. Kingwell, Garth Johnson, Joseph O’Neil, Darren M. Roffey, Vivian Borsella (The Ottawa Hospital); Victoria Avram (Juravinski Hospital and Cancer Centre).

United States: Todd M. Oliver, Vicki Jones, Michelle Vogt (Boone Hospital Center—Columbia Orthopaedic Group); Clifford B. Jones, James R. Ringler, Terrence J. Endres, Debra L. Sietsema, Jane E. Walker (Orthopaedic Associates of Michigan); Kyle J. Jeray, J. Scott Broderick, David R. Goetz, Thomas B. Pace, Thomas M. Schaller, Scott E. Porter, Michael L. Beckish, John D. Adams, Benjamin B. Barden, Aaron T. Creek, Stephen H. Finley, Jonathan L. Foret, Garland K. Gudger Jr., Richard W. Gurich Jr., Austin D. Hill, Steven M. Hollenbeck, Lyle T. Jackson, Kevin K. Kruse, III, Wesley G. Lackey, Justin W. Langan, Julia Lee, Lauren C. Leffler, Timothy J. Miller, R. Lee Murphy, Jr., Lawrence K. O’Malley II, Melissa E. Peters, Dustin M. Price, John A. Tanksley, Jr., Erick T. Torres, Dylan J. Watson, Scott T. Watson, Stephanie L. Tanner, Rebecca G. Snider, Lauren A. Nastoff, Shea A. Bielby, Robert J. Teasdall (Greenville Health System); Julie A. Switzer, Peter A. Cole, Sarah A. Anderson, Paul M. Lafferty, Mengnai Li, Thuan V. Ly, Scott B. Marston, Amy L. Foley, Sandy Vang, David M. Wright (Regions Hospital/University of Minnesota); Andrew J. Marcantonio, Michael S.H. Kain, Richard Iorio, Lawrence M. Specht, John F. Tilzey, Margaret J. Lobo, John S. Garfi (Lahey Hospital & Medical Center); Heather A. Vallier, Andrea Dolenc, Mary Breslin (MetroHealth Medical Center); Michael J. Prayson, Richard Laughlin, L. Joseph Rubino, Jeddah May, Geoffrey Ryan Rieser, Liz Dulaney-Cripe, Chris Gayton (Miami Valley Hospital); James Shaer, Tyson Schrickel, Barbara Hileman (St. Elizabeth Youngstown Hospital); John T. Gorczyca, Jonathan M. Gross, Catherine A. Humphrey, Stephenates, John P. Ketz, Krista Noble, Allison W. McIntyre, Kaili Pecorella (University of Rochester Medical Center); Craig A. Davis, Stuart Weinerman, Peter Weingarten, Philip Stull, Stephen Lindenbaum, Michael Hewitt, John Schwappach, Janell K. Baker, Tom Rutherford, Heike Newman, Shane Lieberman, Erin Finn, Kristin Robbins, Meghan Hurley, Lindsey Lyle, Khalia Mitchell, Kieran Browner, Erica Whatley, Krystal Payton, Christina Reeves (Colorado Orthopedic Consultants); Lisa K. Cannada, David E. Karges, Sarah A. Dawson (St. Louis University Hospital); Samir Mehta, John Esterhai, Jaime Ahn, Derek Donegan, Annamarie D. Horan, Patrick J. Hesketh, Evan R. Bannister (University of Pennsylvania); Jonathan P. Keeve, Christopher G. Anderson, Michael D. McDonald, Jodi M. Hoffman (Northwest Orthopaedic Specialists); Ivan Tarkin, Peter Siska, Gary Gruen, Andrew Evans, Dana J. Farrell, James Irrgang, Arlene Luther (University of Pittsburgh Medical Center); William W. Cross III, Joseph R. Cass, Stephen A. Sems, Michael E. Torchia, Tyson Scrabeck (Mayo Clinic); Mark Jenkins, Jules Dumais, Amanda W. Romero (Texas Tech University Health Sciences Center—Lubbock); Carlos A. Sagebiel, Mark S. Butler, James T. Monica, Patricia Seuffert (University Orthopaedic Associates, LLC); Joseph R. Hsu, Daniel Stinner, James Ficke, Michael Charlton, Matthew Napierala, Mary Fan (US Army Institute of Surgical Research); Paul Tornetta III, Chadi Tannoury, Hope Carlisle, Heather Silva (Boston University Medical Center); Michael Archdeacon, Ryan Finnian, Toan Le, John Wyrick, Shelley Hess (UC Health/University of Cincinnati Medical Center); Michael L. Brennan, Robert Probe, Evelyn Kile, Kelli Mills, Lydia Clipper, Michelle Yu, Katie Erwin (Scott and White Memorial Hospital); Daniel Horwitz, Kent Strohecker, Teresa K. Swenson (Geisinger Medical Center); Andrew H. Schmidt, Jerald R. Westberg (Hennepin County Medical Center); Kamran Aurang, Gary Zohman, Brett Peterson, Roger B. Huff (Kaiser Permanente); Joseph Baele, Timothy Weber, Matt Edison (OrthoIndy Trauma St. Vincent Trauma Center); Jessica Cooper McBeth (Santa Clara Valley Medical Center); Karl Shively, Janes P. Ertl, Brian Mullis, J. Andrew Parr, Ripley Woman, Valda Frizzell, Molly M. Moore (Indiana University—Eskenazi Health Services); Charles J. DePaolo, Rachel Alosky, Leslie E. Shell, Lynne Hampton, Stephanie Shepard, Tracy Nanney, Claudine Cuentz (Mission Hospital Research Institute); Robert V. Cantu, Eric R. Henderson, Linda S. Eickhoff (Dartmouth-Hitchcock Medical Center); E. Mark Hammerberg, Philip Stahel, David Hak, Cyril Mauffrey, Corey Henderson, Hannah Gissel, Douglas Gibula (Denver Health and Hospital Authority); David P. Zamorano, Martin C. Tyran, Deeba Pourmand, Deanna Lawson (University of California Irvine Medical Center); Gregory J. Della Rocca, Brett D. Crist, Yvonne M. Murtha, Linda K. Anderson (University of Missouri Health Care); Colleen Linehan, Lindsey Pilling (Covenant Healthcare of Saginaw); Courtland G. Lewis, Stephanie Caminiti, Raymond J. Sullivan, Elizabeth Roper (Hartford Hospital); William Obremsky, Philip Kregor, Justin E. Richards, Kenya Stringfellow (Vanderbilt University Medical Center); Michael P. Dohm, Abby Zellar (Western Slope Study Group).

The Netherlands: Michiel J.M. Segers, Jacco A.C. Zijl, Bart Verhoeven, Anke B. Smits, Jean Paul M. de Vries, Barbara Fioole, Henk van der Hoeven, Evert B.M. Theunissen, Tammo S. de Vries Reilingh, Lorrieke Goverta, Philippe Wittich, Maurits de Brauw, Jan Wille, Peter M.N.Y.M. Go, Ewan D. Ritchie, Ronald N. Wessel, Eric R. Hammacher (St. Antonius Ziekenhuis); Martin

J. Heetveld, Gijs A. Visser, Heyn Stockmann, Rob Silvis, Jaap P. Snellen, Bram Rijbroek, Joris J.G. Scheepers, Erik G.J. Vermeulen, Michiel P.C. Sireen, Ronald Vuylensteke, Hans L.F. Brom, Herman Rijna (Kennemer Gasthuis); Piet A.R. de Rijcke, Cees L. Koppert, Steven E. Buijk, Richard P.R. Groenendijk, Imro Dawson, Geert W.M. Tetteroo, Milko M.M. Bruijninckx, Pascal G. Doornbosch, Eelco J.R. de Graaf (Utrechtse Ziekenhuis); Maarten van der Elst, Carmen C. van der Pol, Martijne van 't Riet, Tom M. Karsten, Mark de Vries, Niels W.L. Schen, G. Ben Schmidt, W.H. Hoffman (Reinier de Graaf Gasthuis); Maarten P. Simons, Frank H.W.M. van der Heijden, W. Jaap Willems, Frank R.A.J. de Meulemeester, Cor P. van der Hart, Kahn Turckan, Sebastiaan Festen, Frank de Nies, Robert Haverlag, Nico J.M. Out, Jan Bosma (Onze Lieve Vrouwe Gasthuis); Albert van Kampen, Jan Bier, Arie B. van Vugt, Michael J.R. Edwards, Taco J. Blokhuis, Jan Paul M. Frölke, Leo M.G. Geeraedts, Jean W.M. Gardieniers, Edward C.T.H. Tan, Lodewijk M.S.J. Poelhekke, Maarten C. de Waal Malefijt, Bart Schreurs (University Medical Center St. Radboud); Gert R. Roukema, Hong A. Jospautra, Paul Keller, Peter D. de Rooij, Hans Kuiken, Han Boxma, Berry I. Cleffken, Ronald Liem (Maaststad Ziekenhuis); Steven J. Rhemrev, Coks H.R. Bosman, Alexander de Mol van Otterloo, Jochem Hoogendoorn, Alexander C. de Vries, Sven A.G. Meylaerts (Medisch Centrum Haaglanden); Michiel H.J. Verhofstad, Joost Meijer, Teun van Egmond, Frank H.W.M. van der Heijden, Igor van der Brand (St. Elisabeth Ziekenhuis); Peter Patka, Martin G. Eversdijk, Rolf Peters, Dennis Den Hartog, Oscar J.F. Van Wae, Pim Opral (Erasmus MC, University Medical Center Rotterdam); Harm M van der Vis, Martti Campo, Ronald Verhagen, G.H. Robert Albers, Arthur W. Zurcher (Tergooi Ziekenhuis); Rogier K.J. Simmermacher, Jeroen van Mulken, Katrijn van Wessem, Taco J. Blokhuis, Steven M. van Gaalen, Luke P.H. Leenen (University Medical Center Utrecht); Maarten W.G.A. Bronkhorst, Onno R. Guicherit (Bronovo Ziekenhuis); J. Carel Goslings, Robert Haverlag, Kees Jan Ponsen (Academic Medical Center).

International: Mahesh Bhatia, Vinod Arora, Vivek Tyagi (RLB Hospital and Research Centre, India); Susan Liew, Harvinder Bedi, Ashley Carr, Hamish Curry, Andrew Chia, Steve Csongvay, Craig Donohue, Stephen Doig, Elton Edwards, Greg Etherington, Max Esser, Andrew Gong, Arvind Jain, Doug Li, Russell Miller, Ash Moaveni, Matthias Russ, Lu Ton, Otis Wang, Adam Dowrick, Zoe Murdoch, Claire Sage (The Alfred, Australia); Frede Frihagen, John Clarke-Jenssen, Geir Hjorthaug, Torben lanssen, Asgeir Amundsen, Jan Egil Brattgård, Tor Borch, Berthe Bøe, Bernhard Flato, Sondre Hasselund, Knut Jørgen Haug, Kim Hemlock, Tor Magne Høseth, Geir Jomaas, Thomas Kibsgård, Tarjei Lona, Gilbert Moatshe, Oliver Müller, Marius Molund, Tor Nicolaisen, Fredrik Nilsen, Jonas Rydinge, Morten Smedsrød, Are Stødle, Axel Trommer, Stein Ugland, Anders Karsten, Guri Eikås, Elise Berg Vesterhus, Anne Christine Brekke (Oslo University Hospital, Norway); Ajay Gupta, Neeraj Jain, Farah Khan (Nirmal Hospital, India); Ateet Sharma, Amir Sanghavi, Mittal Trivedi (Satellite Orthopaedic Hospital and Research Centre, India); Anil Rai, Subash, Kamal Rai (Highway Hospital, India); Vineet Yadav, Sanjay Singh, Kamal Rai (Popular Hospital, India); Kevin Tetsworth, Geoff Donald, Michael Weinrauch, Paul Pincus, Steven Yang, Brett Halliday, Trevor Gervais, Michael Holt, Annette Flynn (Royal Brisbane and Women's Hospital, Australia); Amal Shankar Prasad, Vimlesh Mishra (Madhuraj Nursing Home, India); D.C. Sundareswaran, Angshuman Khanra (M.S. Ramaiah Medical College & Hospital, India); Joe Joseph Cherian, Davy J Olakengil, Gaurav Sharma (St John's Medical College Hospital, India); Marinis Pirpiris, David Love, Andrew Bucknill, Richard J Farrugia (Royal Melbourne Hospital, Australia); Hans-Christoph Pape, Matthias Knobe, Roman Pfeifer (University of Aachen Medical Center, Germany); Peter Hull, Sophie Lewis, Simone Evans (University Hospitals, England); Rajesh Nanda, Rajanikanth Logishetty, Sanjeev Anand, Carol Bowler (University Hospital of North Tees, England); Akhil Dadi, Naveen Palla, Utsav Ganguly (Sunshine Hospital, India); B. Sachidananda Rai, Janakiraman Rajakumar (Unity Health Complex, India); Andrew Jennings, Graham Chuter, Glynis Rose, Gillian Horner (University Hospital of North Durham and Darlington Memorial Hospital, England); Callum Clark, Kate Eke (Wexham Park Hospital, England); Mike Reed, Dominic Inman, Chris Herriott, Christine Dobb (Northumbria Healthcare NHS Foundation Trust, England).

Sofia Bzovsky, MSc<sup>1,2</sup>  
Nathan N. O'Hara, PhD<sup>3</sup>

Gerard P. Slobogean, MD<sup>4</sup>

Sheila Sprague, PhD<sup>1,5</sup>

Daniel E. Axelrod, MD, MSc<sup>1</sup>

Graeme Hoit, MD<sup>6</sup>

Kiara Pannozzo, MPH<sup>1</sup>

Mohit Bhandari, MD, PhD, FRCSC<sup>1,5</sup>

Marc Swiontkowski, MD<sup>7</sup>

Emil Schemitsch, MD, FRCSC<sup>8</sup>

Rudolf W. Poolman, MD, PhD<sup>2,9</sup>

the FAITH Investigators

<sup>1</sup>Division of Orthopaedic Surgery, Department of Surgery, McMaster University, Hamilton, Ontario, Canada

<sup>2</sup>Department of Orthopaedic Surgery, Leiden University Medical Center, Leiden, Netherlands

<sup>3</sup>R Adams Cowley Shock Trauma Center, Department of Orthopaedics, University of Maryland School of Medicine, Baltimore, Maryland

<sup>4</sup>Department of Orthopaedics, University of California, Irvine, California

<sup>5</sup>Department of Health Research Methods, Evidence, and Impact, McMaster University, Hamilton, Ontario, Canada

<sup>6</sup>Division of Orthopaedic Surgery, University of Toronto, Toronto, Ontario, Canada

<sup>7</sup>Department of Orthopaedic Surgery, University of Minnesota, Minneapolis, Minnesota

<sup>8</sup>Department of Surgery, Western University, London, Ontario, Canada

<sup>9</sup>Department of Orthopaedic Surgery, Joint Research, OLVG, Amsterdam, Netherlands

E-mail address for S. Bzovsky: bzovskys@mcmaster.ca

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