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## Re. Selecting Optimal Subgroups for Treatment Using Many Covariates

### To the Editor:

In a recent publication, VanderWeele et al.<sup>1</sup> considered the task of finding a treatment subgroup that maximizes the mean potential outcome. They showed that the task can sometimes be considerably simplified by deriving optimal treatment assignment rules of a simple form: assign treatment in a greedy fashion to all individuals with the next largest benefit (i.e., the difference in potential outcome means given covariates) or the next highest benefit–cost ratio (with cost being a positive function of baseline covariates) until the resource or cost constraint, respectively, is exceeded. As they state in their eAppendix; <http://links.lww.com/EDE/B655>, the optimality of the rules relies critically on the assumption that there are no ties between individuals. Although tied treatment effects or benefit–cost ratios may occur with many covariates, they are perhaps more realistic when few and only discrete baseline variables are considered to define treatment rules.

Consider for example the setting of the Table and suppose that the total cost may not exceed 130. According to the rule of VanderWeele et al.,<sup>1</sup> individuals in the first stratum should be assigned treatment. Because the presented rules assign treatment to either all or no individuals in any given stratum, no more

**TABLE.** Characteristics of Hypothetical Population of Size 100 with Baseline Covariates Forming Five Strata

	Stratum				
	1	2	3	4	5
Number of individuals	25	20	10	15	30
Conditional mean potential outcome					
Under no treatment	–5	4	0	–5	–5
Under treatment	15	20	20	5	–15
Cost of treatment per individual	4	4	5	10	10
Benefit–cost ratio	5	4	4	1	–1

If those and only those in stratum 1 are treated, the total cost is  $25 \times 4 = 100$  and the mean potential outcome is. If those and only those patients in strata 2 and 3 are treated, the total cost is, and the mean potential outcome is. If patients in stratum 1 are treated with probability 1, patients in strata 2 and 3 with probability 3/13, and the rest with probability 0, the expected total cost is and the mean potential outcome is.

individuals can be selected without violating the cost constraint. This rule yields a mean potential outcome of 2.3. However, because of ties, a better rule that likewise selects either all or no individuals of a stratum, does exist: assign treatment to strata 2 and 3 (with a mean potential outcome of 2.5). Thus, in the presence of ties, the optimal rule need not be greedy (see also the literature on the classic knapsack problem; e.g., Korte and Vygen<sup>2</sup>). We note that a better rule may be obtained by augmenting our data with a sequence of independent, possibly unfair, coin tosses. As shown in the eAppendix; <http://links.lww.com/EDE/B655> (but see also Luedtke and van der Laan<sup>3</sup>), maximizing the mean potential outcome across rules of this kind is achieved in the cost-constrained setting by treating those with a benefit–cost ratio strictly greater than some positive constant and a random selection of those with a benefit–cost ratio that equals that constant. For our example, this means treating all members of stratum 1 as well as those members of strata 2 and 3 whose independent coin toss, with probability 3/13 of showing heads, results in heads (mean potential outcome: 3.5).

It seems unlikely that these treatment rules would be implemented via biased coin tosses in real-world settings. If resources are made available in a single batch, one could calculate the

amount of resources that would need to be allocated to the “always-treat” portion of the population, reserve this portion of resources for always-treat individuals, and then allocate the remainder to the “sometimes-treat” portion of the population on a first-come, first-serve basis until that portion of resources runs out. Bias could however be introduced by doing this, for example, when sometimes-treat individuals who visit the clinic more frequently are systematically less (or more) likely to benefit from treatment. However, there may be ways to account for this (e.g., by including frequency of visits as a covariate).

Finally, we add that with multiple treatment levels and cost constraints, mean potential outcomes need not be optimized by the greedy approach of assigning to subjects the treatment level with the highest benefit–cost ratio above or at treatment level-specific thresholds (to satisfy cost constraints), even if the observed data are augmented with a sequence of independent coin tosses (eAppendix; <http://links.lww.com/EDE/B655>). Regardless of the form the rule should take, however, we encourage researchers to follow VanderWeele et al.<sup>1</sup> in taking a more formal approach to “precision medicine” with clearly specified objectives, so that the optimal rule form may be derived and estimation strategies be evaluated.

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## The Authors Respond

**To the Editor:**

We thank Penning de Vries et al<sup>1</sup> for the illustration of optimal subgroup selection for treatment in the case of ties. In our article,<sup>2</sup> we had addressed this setting of ties by referring the reader to the formal treatment given in Luedtke and van der Laan<sup>3–5</sup> in our eAppendix; (<http://links.lww.com/EDE/B466>). The discussion in our main article was intended to focus on the simpler cases to keep the exposition as accessible as possible, but it is indeed good to have these matters more explicitly discussed in the epidemiologic literature.

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## The Role of Blood Cell Composition in Epidemiologic Studies of Telomeres

**To the Editor:**

Leukocyte telomere length is a widely studied, but inconsistent, marker of disease risk. As telomere length varies

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by blood cell subtype,<sup>1</sup> measurements represent a weighted average across constituent blood cells. Proportions of leukocyte subtypes can differ greatly between individuals and these differences may introduce extraneous inter-individual variation in telomere length estimates. The degree of variation attributed to differences in leukocyte subtype composition is largely unknown, in part because cell sorting methods require fresh blood samples rarely available in epidemiologic studies.

Studies of blood DNA methylation face similar challenges. Like telomeres, blood DNA methylation differs by leukocyte subtype, and there are well-documented examples where failure to adjust for leukocyte proportions demonstrably leads to biased effect estimates.<sup>2</sup> A method widely employed in epidemiologic studies to disentangle leukocyte subtype proportions using patterns of DNA methylation<sup>3</sup> may be useful in research on leukocyte telomere length. Here, we use methylation and telomere length measurements from the same blood DNA samples to examine the effect of leukocyte subtype composition on telomere length measurements.

We used existing data on relative leukocyte telomere length (rLTL)<sup>4</sup> and genome-wide DNA methylation<sup>5</sup> measured in the same blood samples from a subsample of 445 non-Hispanic white women enrolled in the Sister Study (median age, 57; interquartile range, 36–64). rLTL was assessed using multiplex quantitative polymerase chain reaction and standardized as z-scores.<sup>4</sup> The study was approved by the institutional review boards of the National Institute of Environmental Health Sciences and the Copernicus Group. We assessed leukocyte composition by applying a validated deconvolution approach to HumanMethylation450 BeadChip data to estimate proportions of six distinct subtypes (CD8+ and CD4+ T-cells, B-cells, natural killers, monocytes, and granulocytes).<sup>3</sup> We first assessed Pearson correlations between rLTL and each estimated leukocyte proportion. As age is strongly, inversely associated with telomere length,<sup>6</sup> we used the age-telomere length relationship to study the