



Development and validation of a prognostic score to predict mortality in patients with acute-on-chronic liver failure

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Background & Aims: Acute-on-chronic liver failure (ACLF) is a frequent syndrome (30% prevalence), characterized by acute decompensation of cirrhosis, organ failure(s) and high short-term mortality. This study develops and validates a specific prognostic score for ACLF patients.

Methods: Data from 1349 patients included in the CANONIC study were used. First, a simplified organ function scoring system (CLIF Consortium Organ Failure score, CLIF-C OFs) was developed to diagnose ACLF using data from all patients. Subsequently, in 275 patients with ACLF, CLIF-C OFs and two other independent predictors of mortality (age and white blood cell count) were combined to develop a specific prognostic score for ACLF (CLIF

Consortium ACLF score [CLIF-C ACLFs]). A concordance index (C-index) was used to compare the discrimination abilities of CLIF-C ACLF, MELD, MELD-sodium (MELD-Na), and Child-Pugh (CPs) scores. The CLIF-C ACLFs was validated in an external cohort and assessed for sequential use.

Results: The CLIF-C ACLFs showed a significantly higher predictive accuracy than MELDs, MELD-Nas, and CPs, reducing (19–28%) the corresponding prediction error rates at all main time points after ACLF diagnosis (28, 90, 180, and 365 days) in both the CANONIC and the external validation cohort. CLIF-C ACLFs computed at 48 h, 3–7 days, and 8–15 days after ACLF diagnosis predicted the 28-day mortality significantly better than at diagnosis.

Conclusions: The CLIF-C ACLFs at ACLF diagnosis is superior to the MELDs and MELD-Nas in predicting mortality. The CLIF-C ACLFs is a clinically relevant, validated scoring system that can be used sequentially to stratify the risk of mortality in ACLF patients.

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Abbreviations: ACLF, acute-on-chronic liver failure; AD, acute decompensation; CANONIC study, EASL-CLIF Acute on chrONIC liver failure study; CLIF, chronic liver failure; CLIF-C ACLFs, CLIF Consortium ACLF score; CLIF-C OFs, CLIF Consortium organ failure score; CLIF-SOFAs, CLIF-sequential organ failure assessment score; CPs, Child-Pugh score; E, epinephrine; EASL, European Association for the Study of the Liver; FiO₂, fraction of inspired oxygen; HE, hepatic encephalopathy; INR, international normalized ratio; MAP, mean arterial pressure; MELD, model of end-stage liver disease; MELD-Nas, MELD-sodium score; NE, norepinephrine; PaO₂, partial pressure of arterial oxygen; SOFA, sequential organ failure assessment; SpO₂, pulse oximetric saturation.

Introduction

Acute-on-chronic liver failure (ACLF) is a syndrome characterised by acute decompensation of cirrhosis, organ failure(s) and high short-term mortality [1], which was recently defined in the CANONIC study [2]. This was a large prospective observational investigation carried out in 29 European university hospitals. It included 1349 consecutive patients admitted with acute



decompensation of cirrhosis (ascites, bacterial infection, gastro-intestinal haemorrhage, and hepatic encephalopathy) and followed-up for one year. The CANONIC study was organised in the setting of the European Association for the Study of the Liver – Chronic Liver Failure Consortium (EASL CLIF-C).

The results of the CANONIC study showed that ACLF occurs most frequently in relatively young individuals, affects approximately 30% of hospitalised cirrhotic patients, may develop in patients without previous decompensation, is associated with a 28-day mortality rate of 33% (51% at 90 days) and is distinct from 'mere' decompensation of cirrhosis. ACLF is the most frequent indication for admission to the intensive care unit (ICU) [3]. In the US, about 200,000 patients with cirrhosis are hospitalised each year of which about 26,000 patients require ICU care [1,3,4]. An average ICU admission costs about \$116,200 and costs for the health care system are \$3 billion for cirrhotic patients requiring intensive care [3].

The diagnostic criteria for ACLF in the CANONIC study were based on the Chronic Liver Failure-SOFA score (CLIF-SOFAs), an adaptation for cirrhotic patients of the sepsis organ failure assessment score (SOFAs) widely used in the ICU [4]. The CLIF-SOFAs, however, is complex (based on 6 subscores, each with a 5-point range, assessing liver, kidney, brain, coagulation, respiration, and circulation), is based on consensus and expert opinion rather than data, and does not significantly improve the prediction accuracy of the model for end-stage liver disease (MELD) and MELD-sodium (MELD-Na) scores [5–6].

The current study was aimed to simplify the original CLIF-SOFAs and develop a new score for ACLF patients (CLIF Consortium ACLF score, CLIF-C ACLFs) with a higher prognostic accuracy than the CLIF-SOFA, MELD, MELD-Na, and the Child-Pugh (CP) scores [7] for patients with ACLF. The study therefore had four main objectives. First, to develop a simpler and validated organ failure score (CLIF Consortium Organ Function score, CLIF-C OFs) for the diagnosis and grading of ACLF. Second, to design a more accurate prognostic score for ACLF patients (CLIF-C ACLFs), using the CLIF-C OFs and other prognostic clinical and biochemical data. Third, to compare the prognostic accuracy of the CLIF-C ACLFs to that of MELDs, MELD-Nas, and CPs. Fourth, to validate the prognostic accuracy of the CLIF-C ACLFs in an external prospective cohort of consecutive patients hospitalized in a single ICU and assess the score for sequential use.

The CANONIC study database was used as a derivation set for several reasons: first, it includes a large series of patients with acute decompensated cirrhosis and also with ACLF. Second, CANONIC patients were closely and prospectively followed-up for up to 1-year. Third, the population with ACLF included patients developing the syndrome either at study inclusion or during the hospitalization period. Finally, as patients were recruited from 29 centres in Europe, the CANONIC data are representative of the European patient population.

Patient and methods

Study populations

The study was performed in patients from three different populations. Both the derivation and the validation datasets came from studies approved by Ethical Review Boards of all study sites.

1. The CLIF-C OFs was developed using the baseline (i.e., enrolment) data of the whole CANONIC study population, which included 1349 patients (out of 2149 screened) admitted to 29-European hospitals within a period of 6 months for the treatment of decompensated cirrhosis. These patients were prospectively followed-up for one year [2]. In most patients (52%) the aetiology of cirrhosis was alcoholic, in 19.5% it was associated with chronic hepatitis C virus infection, and in 9.6% it was due to both alcohol and hepatitis C. In the remaining 18.9%, cirrhosis was due to other causes. 345 patients (26.8%) had no history of previous decompensation. Causes of hospitalization at study enrolment were ascites (66.8%), hepatic encephalopathy (34.3%), bacterial infections (24.2%), and/or gastro-intestinal haemorrhage (16.4%). 196 patients (14.6%) were admitted to the ICU. The MELDs at enrolment in the whole series was 18.8 (SD: 7.5), and CPs was 9.7 (SD: 2.1).
2. The CLIF-C ACLFs was developed using data from 275 CANONIC patients with ACLF at enrolment, or those who developed ACLF within 28-days post-enrolment. Diagnostic criteria for organ failures in the CANONIC study are described in [Supplementary Table 1](#). The diagnosis of ACLF was based on the presence of at least renal failure or any other single organ failure if associated with renal dysfunction (serum creatinine 1.5–1.9 mg/dl) and/or grade I–II hepatic encephalopathy (ACLF-1). Patients with two organ failures were graded as ACLF-2 and those with three or more organ failures as ACLF-3. At study enrolment, 70.1% of these patients were admitted to the ICU.
3. The external validation of the CLIF-C ACLFs was carried out using data from 225 ACLF patients consecutively admitted to the ICU at the Paul Brousse hospital, Villejuif, France [8]. Despite differences in the proportion of patients admitted to the ICU, this series of patients was selected as the validation set for the following reasons: (1) it was a prospective cohort; (2) all patients had the data needed for ACLF diagnosis and score calculations; (3) patients were followed-up for 90-days. [Table 1](#) shows the clinical characteristics of patients included in the derivation and validation sets.

Study outcomes

The main study outcomes included all-cause mortality at 28, 90, 180, and 365 days after enrolment. All CANONIC patients were closely followed-up during the first 28 days. Subsequently, data on vital signs, causes of death and liver transplantation were obtained 3, 6, and 12 months after enrolment. The same information was collected for all ACLF patients in the validation cohort, who were followed-up for 3 months only. In both derivation and validation sets, there were no losses due to follow-up. At 90 days, 38/275 (13.8%) ACLF patients underwent liver transplantation in the CANONIC dataset and 22/225 (9.8%) in the validation cohort. One year after study enrolment, 53/275 (20%) CANONIC patients with ACLF had been transplanted.

All the data required to compute CLIF-C ACLFs (as well as those used to compute MELDs, MELD-Nas, and CPs) were measured at the time of ACLF diagnosis (either at enrolment or within the 28-day post-enrolment follow-up). Patients' parameters were collected based on standard lab measurements performed at study enrolment and at day 2, 3–7, 8–14, 15–21, and 22–28 during the hospitalization. A central laboratory was not used for sample analyses. However, to assure the comparability of lab results, site labs were requested to use the same units and normal ranges; extensive remote monitoring and quality control were carried out during and after study termination.

Statistical methods

Variables used for the CLIF-C OFs and CLIF-C ACLFs were measured at enrolment and at ACLF diagnosis. The original CLIF-SOFAs included 6 subscores – one for each organ/system – each of them ranging from 0 to 4 ([Supplementary Table 1](#)). The 5 categories included in CLIF-SOFAs subscores and the corresponding cut-off values were derived from a consensus [2]. In the current study we assessed whether the cut-offs could be modified and/or if the number of the original categories of each subscore could be reduced maintaining the predictive ability of the aggregated score.

In univariate statistical comparisons, the χ^2 test was used for categorical variables, Student's *t* test and Mann-Whitney test for continuous variables. McNemar's test and paired *t* test were used to compare repeated measurements of categorical and continuous variables, respectively. Proportional hazards models considering liver transplantation as a competing risk (PH-CR) were used to identify additional independent factors of mortality not included in the CLIF-C OFs system. Transplanted patients were considered as censored and the survival function was adjusted for the risk of liver transplantation at each study time point [9,10].

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Table 1. Baseline characteristics and outcome of patients with an acute-on-chronic liver failure (ACLF) episode included in the CANONIC study and in the external validation datasets.

Patients' characteristics	Derivation set (CANONIC patients) (N = 275)	External validation set (N = 225)	p value
Age (yr)	54.5 ± 12.1	55.1 ± 11.1	0.567
Male sex	176 (64.0)	171 (76.0)	0.005
Presence of ascites	218 (80.2)	206 (93.2)	<0.001
Etiology of cirrhosis			
Alcohol	147 (54.7)	158 (70.2)	<0.001
HCV	40 (14.9)	24 (10.7)	0.218
Alcohol + HCV	29 (10.8)	14 (6.2)	0.100
Site of hospitalization at study enrolment			
Intensive care unit	192 (70.1)	225 (100)	<0.001
Data used to compute CLIF-OF score			
Serum bilirubin (mg/dl)	12.3 ± 11.5	13.3 ± 12.7	0.360
Serum creatinine (mg/dl)	2.0 ± 1.3	2.0 ± 1.5	0.721
HE grade I-II	102 (37.5)	111 (49.3)	<0.001
HE grade III-IV	60 (22.1)	62 (27.6)	
INR	2.2 ± 1.0	3.8 ± 2.8	<0.001
Serum sodium (mEq/L)	135 ± 7	134 ± 8	0.137
White-cell count (x10 ⁹ cells/L)	10.0 ± 6.3	7.8 ± 4.3	<0.001
Use of vasopressors	108 (39.7)	143 (63.6)	<0.001
PaO ₂ /FiO ₂ 200-300 or SpO ₂ /FiO ₂ 214-357	71 (25.8)	93 (41.3)	<0.001
PaO ₂ /FiO ₂ ≤200 or SpO ₂ /FiO ₂ ≤214	39 (14.2)	86 (38.2)	
Acute-on-chronic liver failure (ACLF) grade and scores			
ACLF grade 1	121 (44.0)	51 (22.6)	
ACLF grade 2	110 (40.0)	71 (31.6)	<0.001
ACLF grade 3	44 (16.0)	103 (45.8)	
Child-Pugh score	11.1 ± 2.0	11.4 ± 1.9	0.104
MELD score	28 ± 8	31 ± 8	<0.001
MELD-Na score	30 ± 7	33 ± 7	<0.001
Mortality rates			
28-Day mortality	93 (33.8)	117 (52.0)	<0.001
90-Day mortality	133 (48.4)	141 (62.7)	0.002
6-Month mortality	143 (52.0)		
1-Year mortality	159 (57.8)		

Data are numbers of patients (%) or mean ± SD.

HE, hepatic encephalopathy; FiO₂, fraction of inspired oxygen; PaO₂, partial pressure of arterial oxygen; SpO₂, pulse oximetric saturation.

Baseline factors, not accounted for in the CLIF-C OFs, and significantly ($p < 0.05$) associated with mortality at 28, 90, 180, and 365 days, were selected for the final models. PH-CR models including the CLIF-C OFs and all the other selected factors were fitted applying a forward step-wise selection method with p -in = 0.05 and p -out = 0.1. The factors that were independently associated with mortality at the main time points were included in the CLIF-C ACLFs. The coefficients estimated for each factor in the 28-day model, which provided the best predictive ability, were used as relative weights to compute the CLIF-C ACLFs.

The calibration of the CLIF-C ACLFs was assessed by comparing the actual observed risk and the average probability of dying at different time points predicted by the score. The Hosmer-Lemeshow test was used to assess the corresponding goodness-of-fit. The Harrell's concordance index (C-index) was used to assess the score's discrimination ability [11,12]. Since the CLIF-C ACLFs was derived based on a PH-CR model, C-index values and the corresponding 95% confidence intervals (CIs) were estimated for each main study time point, treating the transplanted patients as censored at the end of the period, assuming that

none of them would die before [9]. Statistical comparisons of the C-index between the CLIF-C ACLFs, CLIF-SOFAs, CLIF-C OFs, MELDs, MELD-Nas, and CPs were carried out for the main study time points using the integrated discriminating improvement (IDI) statistic [13].

For external validation [14], the CLIF-C ACLFs was computed with the validation data, and score performance was assessed and compared to that of the MELDs, MELD-Nas, and CPs by means of the same methods applied to the CANONIC data. To corroborate the results observed in the derivation and validation sets, a confirmatory analysis was carried out by estimating the Area Under the ROC curve (AUROC) of the CLIF-C ACLFs for predicting 28-day and 90-day mortality. This also allowed us to identify the best score cut-points, which maximized the sensitivity, specificity and predictive values.

The prognostic ability of the CLIF-C ACLFs, and MELDs for sequential use was assessed using a subset of CANONIC patients with ACLF and data available at the time of diagnosis and at 48 h, 3–7 days, and at 8–15 days after diagnosis. Post-diagnosis C-indexes were compared with the baseline by means of paired t tests.

Table 2. The CLIF-organ failure score system.

Organ/system	Subscore = 1	Subscore = 2	Subscore = 3
Liver	Bilirubin <6 mg/dl	Bilirubin ≥6 mg/dl and <12 mg/dl	Bilirubin ≥12 mg/dl
Kidney	Creatinine <2 mg/dl	Creatinine ≥2 mg/dl and <3.5 mg/dl	Creatinine ≥3.5 mg/dl or renal replacement
Brain (West-Haven grade for HE*)	Grade 0	Grade 1-2	Grade 3-4**
Coagulation	INR <2.0	INR ≥2.0 and <2.5	INR ≥2.5
Circulatory	MAP ≥70 mmHg	MAP <70 mmHg	Use of vasopressors
Respiratory			
PaO ₂ /FiO ₂	>300	≤300 and >200	≤200 [#]
or		or	
SpO ₂ /FiO ₂	>357	>214 and ≤357	≤214 [#]

The shaded area describes criteria for diagnosing organ failures.

*HE, hepatic encephalopathy; FiO₂, fraction of inspired oxygen; PaO₂, partial pressure of arterial oxygen; SpO₂, pulse oximetric saturation.

**Patients submitted to Mechanical Ventilation (MV) due to HE and not due to a respiratory failure were considered as presenting a cerebral failure (cerebral subscore = 3).

[#]Other patients enrolled in the study with MV were considered as presenting a respiratory failure (respiratory subscore = 3).

Results

ACLF study populations

Table 1 shows the two ALCF study populations. ALCF patients in the external validation cohort were more frequently male and alcoholic, had more severe ALCF and higher rates of ascites, hepatic encephalopathy, respiratory and circulatory failure than ALCF patients in the CANONIC cohort. Consequently, the MELDs and MELD-Nas and the 28-day and 90-day mortality rates were higher in the validation cohort.

Development of the CLIF-C OFs

For each organ system, two new cut-points were chosen to distinguish three clinical severity categories that were directly correlated with the risk of dying at 28-days. The new cut-off values maximised the ability of the aggregated score (ranging 6 to 18) to predict 28-day mortality. Derivations of the resulting organ system subscores are reported in Table 2. The performance of the CLIF-C OFs (C-index: 0.72) in predicting 28-day mortality was identical to that of the original CLIF-SOFAs (C-index estimate: 0.72 [p = 0.856]) and slightly but significantly superior to those of MELDs, MELD-Nas and CPs (C-index: 0.69 [p = 0.015], 0.68 [p = 0.019] and 0.67 [p = 0.014] respectively).

Development of the CLIF-C ACLFs

Baseline factors, not included in the CLIF-C OFs and significantly associated with short- and long-term mortality, were alcoholic aetiology of cirrhosis, ascites (as clinically diagnosed), AST, serum sodium, serum potassium and white blood cell (WBC) count (Supplementary Table 2). Age was included in all models as a well-known potential confounder. Together with the CLIF-C OFs, age and log-transformed white blood cell count were selected as the best predictors. The CLIF-C ACLFs was computed by applying model coefficients and was trimmed between 0 and 100, since upper values did not modify the expected probabilities of dying by more than 1%:

$$CLIF-C\ ACLFs = 10 \times [0.33 \times CLIF-OFs + 0.04 \times Age + 0.63 \times \ln(WBC\ count) - 2]$$

The probability of death at time “t” can be estimated by the equation:

$$P = 1 - e^{-CI(t) \times \exp(\beta(t) \times CLIF-C\ ACLFs)}$$

CI(t) and β(t) are the cumulated baseline hazard and the score coefficient estimated by the model fitted for time t. At the main time points they are: CI(28) = 0.0022, β(28) = 0.0995; CI(90) = 0.0079, β(90) = 0.0869; CI(180) = 0.0115, β(180) = 0.0824; CI(365) = 0.0231, β(365) = 0.0731.

An online application to estimate the predicted death rate at time t based on the CLIF-C ACLFs will be available at the CLIF Consortium website: <http://www.clifconsortium.com/>. A freely downloadable application is also available (CLIFC ACLFs).

Calibration of the CLIF-C ACLFs

Fig. 1 shows the observed and predicted probability of dying at 28 days by the CLIF-C ACLFs quintiles in the CANONIC and in the validation datasets. In the derivation set, predicted and observed probabilities of death were similar across the quintiles of CLIF-C ACLFs (Hosmer-Lemeshow χ² = 7.9, p = 0.44) and at the four study time points: 28-days: overall predicted 0.28 vs. overall observed 0.34; 90-days: 0.46 vs. 0.48; 6-month: 0.51 vs. 0.52; 1-year: 0.59 vs. 0.58. The calibration of the CLIF-C ACLFs was also assessed for the validation dataset with similar results: observed and estimated probabilities of death were comparable at the main study time points (28-days: predicted 0.45 vs. observed 0.52; 90-days: 0.64 vs. 0.63) and the Hosmer-Lemeshow test for the 28-day mortality did not show a significant lack of fit (χ² = 4.0, p = 0.26).

Discrimination ability of the CLIF-C ACLFs. Comparison with the CLIF-C OFs, CLIF-SOFAs, MELDs, MELD-Nas and the CPs

The C-index of the CLIF-C ACLFs for 28-day, 90-day, 6-month and 1-year mortality (0.76, 0.73, 0.72, and 0.71) was significantly better than those corresponding to the CLIF-C OFs (0.72 [p < 0.001], 0.68 [p < 0.001], 0.67 [p < 0.001], and 0.66 [p = 0.003]), and CLIF-SOFAs (0.72 [p < 0.001], 0.68 [p < 0.001], 0.67 [p < 0.001], and 0.66 [p = 0.002]).

The CLIF-C ACLFs showed a significantly higher predictive discrimination than the MELDs, MELD-Nas, and CPs at 28-days and at the rest of the study time points (Table 3). The

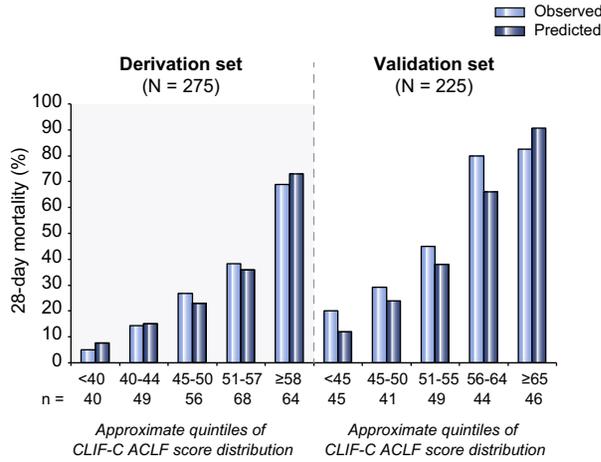


Fig. 1. Observed (light blue) vs. predicted (dark blue) 28-day mortality rates: according to the approximate quintiles of the CLIF-C ACLF score in the ACLF patients included in the derivation set (left side) and validation set (right side). The mortality probabilities predicted using the CLIF-C ACLF score were similar to those observed in both sets of patients, thus indicating a good performance of the score throughout the whole range of CLIF-C ACLF values.

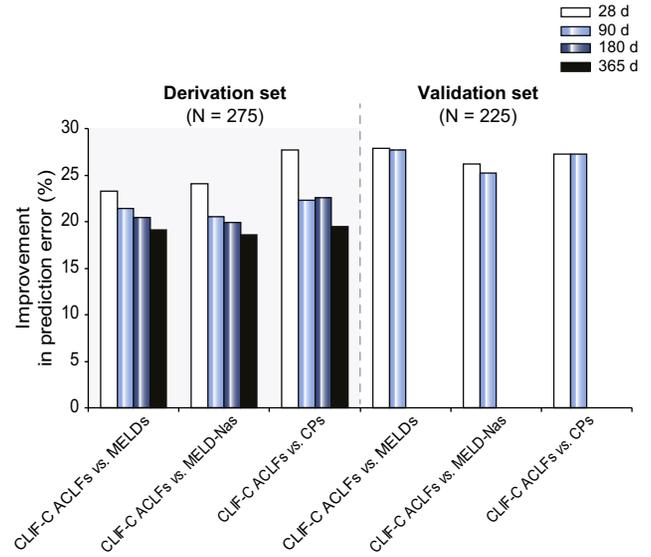


Fig. 2. Percent reduction in prediction error rates of the CLIF-C ACLF score as compared to the Child-Pugh, MELD, and MELD-Na scores at the main study time points. The prediction error rates observed using CLIF-C ACLFs were 19% to 28% lower than those observed for MELDs, MELD-Nas or CPs in the derivation set and 25% to 28% lower in the validation cohort.

absolute improvements of about 7–8 points in the C-index values with respect to the MELDs, MELD-Nas and CPs were consistently significant at all time points. Fig. 2 shows the corresponding percent improvement obtained with the CLIF-C ACLFs in prediction error rate with respect to the other scores (computed as percent reduction in discordance rate of the CLIF-C ACLFs vs. the reference (REF) score, i.e., $100 \times [C\text{-index}_{\text{CLIF-C ACLFs}} - C\text{-index}_{\text{REF}}] / [1 - C\text{-index}_{\text{REF}}]$). The CLIF-C ACLFs consistently improved the prediction error rates observed for the MELDs, MELD-Nas, and CPs by 19%–28% at all study time points.

Validation of the CLIF-C ACLFs

The comparative C-index estimates for the validation data cohort are shown in Table 3. The predictive ability of the CLIF-C ACLFs at

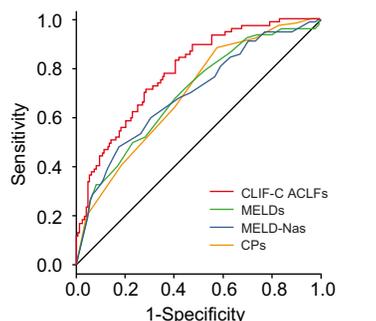
each main time point was significantly better than for the MELDs, MELD-Nas and CPs (Supplementary Fig. 1). The CLIF-C ACLFs improved the 28-day and 90-day mortality predictions by about 25% to 28% as compared to the MELDs, MELD-Nas, and CPs (Fig. 2).

Figs. 3 and 4 show the results of the confirmatory analysis of the predictive ability of the CLIF-C ACLF score for the 28-day and 90-day mortality in the CANONIC patients with ACLF. As compared to the MELDs, MELD-Nas, and CPs, the AUROCs estimated for the CLIF-C ACLFs were significantly higher and indicated a 7% to 11% improvement in the discrimination ability, so confirming the concordance index estimates shown in Table 3. A CLIF-C ACLFs of 40 or lower had a 90% negative predictive value and

Table 3. Predictive discrimination ability of the CLIF-C ACLF score as compared with the MELDs and MELD-Nas. ACLF patients in the CANONIC database and in the validation dataset.

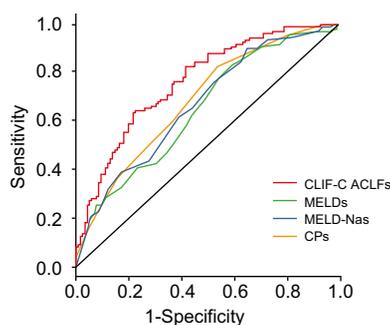
	CLIF-C ACLF score C-index (95% CI)	Child-Pugh score C-index (95% CI)	MELD score C-index (95% CI)	MELD-Na score C-index (95% CI)
CANONIC PATIENTS (N = 275)				
28-day mortality	0.760 (0.715-0.805)	0.668 (0.610-0.726)	0.687 (0.635-0.738)	0.684 (0.632-0.736)
p value vs. CLIF-C*		<0.001	<0.001	<0.001
90-day mortality	0.732 (0.691-0.773)	0.655 (0.605-0.705)	0.659 (0.615-0.710)	0.663 (0.617-0.709)
p value vs. CLIF-C*		<0.001	<0.001	0.001
180-day mortality	0.723 (0.683-0.763)	0.642 (0.593-0.691)	0.652 (0.607-0.697)	0.654 (0.609-0.699)
p value vs. CLIF-C*		<0.001	<0.001	0.001
365-day mortality	0.707 (0.668-0.746)	0.636 (0.588-0.683)	0.638 (0.595-0.682)	0.640 (0.597-0.683)
p value vs. CLIF-C*		<0.001	<0.001	<0.001
Validation database (N = 225)				
28-day mortality	0.744 (0.702-0.787)	0.653 (0.603-0.704)	0.645 (0.593-0.697)	0.648 (0.597-0.700)
p value vs. CLIF-C*		<0.001	<0.001	<0.001
90-day mortality	0.736 (0.696-0.776)	0.647 (0.599-0.695)	0.635 (0.585-0.684)	0.637 (0.588-0.686)
p value vs. CLIF-C*		<0.001	<0.001	<0.001

*p values from the Integrated Discriminating Improvement (IDI) statistics test.



	AUROC (95% CI)	p value vs. CLIF-C ACLF
CLIF-C ACLFs	0.79 (0.73-0.85)	
MELDs	0.70 (0.62-0.77)	0.0089
MELD-Nas	0.70 (0.62-0.77)	0.0097
CPs	0.70 (0.63-0.77)	0.0075

Fig. 3. Accuracy of the CLIF-C ACLFs as compared to that of MELDs, MELD-Nas and CPs in predicting 28-day mortality of the CANONIC patients with ACLF. Comparison of the area under the ROC curves (AUROCs) estimated for each score. The CLIF-C ACLFs showed a significantly higher predictive ability in comparison with all the other scores. (This figure appears in colour on the web.)



	AUROC (95% CI)	p value vs. CLIF-C ACLF
CLIF-C ACLFs	0.76 (0.70-0.83)	
MELDs	0.65 (0.58-0.72)	0.0014
MELD-Nas	0.67 (0.60-0.74)	0.0082
CPs	0.69 (0.62-0.75)	0.0301

Fig. 4. Accuracy of the CLIF-C ACLFs as compared to that of MELDs, MELD-Nas, and CPs in predicting 90-day mortality of the CANONIC patients with ACLF. Comparison of the area under the ROC curves (AUROCs) estimated for each score. The CLIF-C ACLFs showed a significantly higher predictive ability in comparison with all the other scores. (This figure appears in colour on the web.)

97% sensitivity, while a score of 60 or higher allowed for a 82% positive predictive value and 94% specificity. A score cut-off of 51 maximized both the sensitivity (64%) and the specificity (75%), but was found to have lower negative (69%) and positive (70%) predictive values. In the validation set comparable results were observed: a CLIF-C ACLFs of 40 or lower had a 77% negative predictive value and 97% sensitivity, while a score of 60 or higher showed a 91% positive predictive value and 93% specificity.

Sequential use of the CLIF-C ACLFs

Table 4 reports the C-index estimates for the CLIF-C ACLFs and the MELDs computed at ACLF diagnosis and at 48 h, 3–7 days, and 8–15 days after ACLF diagnosis in the 256 CANONIC study

patients with follow-up clinical and laboratory data. Both scores, when used sequentially, improved their predictive performance. The CLIF-C ACLFs computed at 3–7 days and 8–15 days after ACLF diagnosis predicted 28-day and 90-day mortality significantly better than the CLIF-C ACLFs at ACLF diagnosis. Interestingly, the sequence of C-index estimates reflected a consistent difference of 7–8 points between the CLIF-C ACLFs and the MELDs in the corresponding discrimination abilities.

Discussion

The observation in the CANONIC study that ACLF occurs in 30% of hospitalised cirrhotic patients, is associated with a 28-day mortality rate of 32%, has well-defined diagnostic criteria, and is pathophysiologically related to systemic inflammation provides the rationale for new investigations on the mechanism and management of this syndrome [1,2,8,15,16]. Consequently, it is necessary to develop an accurate prognostic score for ACLF patients.

Any specific score for ACLF should have the following characteristics. First, the score should be as simple as possible to facilitate clinical use and be based on the function of the vital organ-systems defining the ACLF syndrome [2]. Second, parameters estimating organ failure should be easy to obtain in standard hospitalization settings [3,8,15,16]. Third, the cut-off levels for the definition of organ system failure should have prognostic significance. Fourth, the score should provide computed figures, estimating risk of mortality but also allow an easy stratification of patients with clinical and prognostic significance. Finally, it should clearly improve the predictive ability of the main prognostic scores currently available (MELDs and CPs).

The first step of this study was to develop a simplified score based on organ function (the CLIF-C OFs), using the data obtained from all patients included in the CANONIC study [2]. As indicated above, the main objective of this new score was to diagnose and grade the severity of ACLF. The CLIF-C OFs, which has a 3-point range per organ system and an aggregated score ranging from 6 to 18, is simpler than the original CLIF-SOFAs used in the CANONIC study from which it derives (5 point ranges and aggregated score range 0–24) and accomplishes all the above mentioned criteria except the last one. Interestingly, the prognostic significance of cut-off levels for the diagnosis of organ failures of the original CLIF-SOFAs were confirmed in the CLIF-C OFs [2]. Therefore, diagnosis of organ failure and the stratification of ACLF proposed by the CANONIC study are also applicable using the CLIF-C OFs. In the CANONIC patients with ACLF, the predictive accuracy of the CLIF-C OFs was identical to that of the CLIF-SOFA score and slightly but significantly better than those of the MELDs, MELD-Nas, and CPs.

The second step of the study was to develop a prognostic score better than the CLIF-C OFs for patients with ACLF. The strategy used was to analyse patients with ACLF in the CANONIC database looking for independent factors associated with mortality not included in the CLIF-OFs and to subsequently fit a final survival model with the CLIF-OFs and the selected factors. Age and white blood cell count (a crude marker of systemic inflammation) were found to be the best predictors. These parameters have also been found to be independently associated with mortality in previous studies [2,8,15–19]. Although in the CANONIC study, white blood cell count was higher in patients with alcoholic liver

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Table 4. Sequential use of the CLIF-C ACLF and MELD score to predict mortality in CANONIC patients with ACLF and post-enrolment clinical and laboratory follow-up.

	28-day mortality				90-day mortality			
	CLIF-C ACLF score		MELD score		CLIF-C ACLF score		MELD score	
	C-index (95% CI)	<i>p</i> value vs. baseline	C-index (95% CI)	<i>p</i> value vs. baseline	C-index (95% CI)	<i>p</i> value vs. baseline	C-index (95% CI)	<i>p</i> value vs. baseline
CLIF-C-ACLF score at enrolment (N = 256)	0.751 (0.701-0.800)		0.679 (0.621-0.737)		0.712 (0.666-0.759)		0.653 (0.602-0.703)	
CLIF-C-ACLF score at 48 hours (N = 186)	0.801 (0.747-0.854)	0.0956	0.721 (0.658-0.783)	0.1890	0.751 (0.700-0.802)	0.1336	0.680 (0.625-0.736)	0.3312
CLIF-C-ACLF score at 3-7 days (N = 189)	0.822 (0.767-0.877)	0.0179	0.749 (0.682-0.815)	0.0389	0.774 (0.722-0.827)	0.0217	0.706 (0.646-0.765)	0.0824
CLIF-C-ACLF score at 8-15 days (N = 154)	0.866 (0.809-0.923)	0.0001	0.799 (0.729-0.870)	0.0008	0.790 (0.733-0.847)	0.0072	0.710 (0.643-0.776)	0.0958

**p* value vs. CLIF-C score at enrolment.

disease and active alcoholism, the data also showed that irrespective of aetiology, higher white blood cell count was associated with more severe grades of ACLF [2]. Through the final model including a CLIF-C OFs with these two new variables, we were able to obtain a new score (CLIF-C ACLFs) ranging between 0 and 100 points, which predicted the risk of mortality accurately.

The concordance index analysis in the derivation cohort showed that the CLIF-C ACLFs was more accurate in predicting short-term and long-term mortality than their predecessors CLIF-SOFAs and CLIF-C OFs (data not shown) and also than MELDs, MELD-Nas, and CPs. The differences in predictive discrimination between the CLIF-C ACLFs and MELDs, MELD-Nas, and CPs at 28 and 90-days were even more significant in the validation cohort. The fact that these patients were more severely ill than those in the derivation cohort probably accounts for this observation. The analysis of the prediction errors, observed with the CLIF-C ACLFs compared with MELDs, MELD-Nas, and CPs in both the derivation and the validation cohorts, is additional evidence of the superiority of the CLIF-C ACLFs over the other scores in predicting mortality in ACLF patients. Using the CLIF-C ACLFs leads to a substantial improvement (19% to 28%) in the discrimination ability observed with MELDs, MELD-Nas, and CPs.

We also assessed whether the CLIF-C ACLF score could be useful if applied sequentially to identify response to an intervention, deterioration despite intervention and perhaps act as a guide to determine whether further interventions are likely to be futile. The data showed that the performance of the CLIF-C ACLFs improved significantly if repeated 48 h, 3–7 d or 8–15 d after the initial determination. These data suggest that, in clinical practice, the CLIF-C ACLFs, like the SOFAs, which is widely used in the ICU for non-cirrhotic patients, can be updated on a daily basis providing additional prognostic information [4]. Currently, no validated evidence-based tools guide the decision-making. Sequential measurements of the CLIF-C ACLFs may help to define potential futility.

In summary, using the CANONIC database we derived a new, evidence-based and simpler organ-failure score (CLIF-C OFs) to diagnose organ failures and ACLF in cirrhotic patients. The CLIF-C OFs includes parameters most commonly used in the management of ACLF patients and shows a prognostic accuracy similar to its predecessor, the CLIF-SOFAs but significantly higher than MELDs, MELD-Nas, and CPs. Combining the CLIF-C OFs with age and white blood cell count, the CLIF-C ACLFs was obtained as a specific prognostic score for ACLF patients and validated in an external cohort of ACLF patients. CLIF-C ACLFs allowed a significant improvement of the discrimination ability as compared to

the MELDs, MELD-Nas, and CPs, as indicated by a 19% to 28% reduction in percent prediction errors observed in both the derivation and validation datasets. The CLIF-C ACLFs also proved to be potentially useful for sequential use after ACLF diagnosis but this will require further validation.

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Conflict of interest

The EASL-CLIF Consortium

The EASL-CLIF Consortium is a network of 63 European university hospitals, aimed at stimulating research on pathophysiology, diagnostic and treatment on Chronic Liver Failure. During the period 2009–2012 the EASL-CLIF Consortium had received unrestricted grants from Grifols and Gambro. Grifols has prolonged its unrestricted grant for an additional period of four years. There is no other support for the Consortium. The Fundació Clinic, a foundation ruled by the Hospital Clinic and University of Barcelona, administers the EASL-CLIF Consortium grants. V. Arroyo (Chairman), M. Bernardi (Vice-Chairman), and members of the Steering Committee have no relationship with Grifols or Gambro other than conferences at international meetings (from which they may receive an honorarium) or as investigators on specific projects unrelated to the consortium. Up to now the EASL-CLIF Consortium has not performed any study promoted by pharmaceutical companies. The scientific agenda of the EASL-CLIF Consortium and the specific research protocols are made exclusively by the Steering Committee members without any participation of pharmaceutical companies.

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R. Jalan received research funding from Vital Therapies, has served on the Scientific Advisory Board for Conatus Pharma,

received lecture fees from Gambro, has on-going research collaboration with Gambro, Grifols, and is the Principal Investigator of an industry sponsored study (Sequana Medical). He is also the inventor of a drug, L-ornithine phenylacetate, which UCL has licensed to Ocera Therapeutics. F. Saliba has received speaker honorarium and/or grant research from Novartis, Astellas, Roche, Merck Sharp & Dohme, Pfizer, Gambro and Vital Therapies. P. Ginès has received speaker honorarium and research funding from Grifols, served on the scientific advisory board for Ferring and Sequana and received research funding from Sequana. J. Cordoba has served as a consultant to Ocera. A. Gerbes has served as a consultant to CSL Behring. S. Zeuzem has served as a consultant to Abbott, Achillion, AstraZeneca, BMS, Boehringer Ingelheim, Gilead, Janssen Cilag, Merck, Novartis, Presidio, Roche, Santaris, Vertex. V. Arroyo has received grant and research support from Grifols. All other authors declare that they have no conflicts of interest.

Authors' contributions

R. Jalan, R. Moreau, M. Pavesi, A. Amoros, F. Saliba, M. Bernardi, and V. Arroyo participated in the writing group. R. Moreau, R. Jalan, P. Ginès, M. Pavesi, F. Durand, T. Gustot, J. Trebicka, and V. Arroyo designed the CANONIC study with input from all other authors. R. Moreau, R. Jalan, M. Pavesi, A. Amoros, F. Saliba, and V. Arroyo participated in data analysis and interpretation. F. Durand, P. Angeli, P. Caraceni, C. Hopf, C. Alessandria, E. Rodríguez, P. Solis-Muñoz, W. Laleman, J. Trebicka, S. Zeuzem, T. Gustot, R. Mookerjee, L. Elkrief, G. Soriano, J. Cordoba, F. Morando, A. Gerbes, B. Agarwal, and D. Samuel participated in patient recruitment and data collection along with investigators of the CANONIC study. F. Saliba, E. Levesque, and D. Samuel provided the data from Paul-Brousse Hospital for external validation. V. Arroyo was responsible for obtaining funding and overall project collaboration.

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Research Article

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Supplementary data

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References

- [1] Jalan R, Gines P, Olson JC, et al. Acute-on chronic liver failure. *J Hepatol* 2012;57:1336–1348.
- [2] Moreau R, Jalan R, Gines P, et al. Acute-on-chronic liver failure is a distinct syndrome developing in patients with acute decompensation in cirrhosis. *Gastroenterology* 2013;144:1426–1437.
- [3] Olson JC, Wendon JA, Kramer DJ, et al. Intensive care of the patient with cirrhosis. *Hepatology* 2011;54:1864–1872.
- [4] Ferreira FL, Bota DP, Bross A, Melot C, Vincent JL. Serial evaluation of the SOFA score to predict outcome in critically ill patients. *JAMA* 2001;286:1754–1758.
- [5] Kamath PS, Kim WR. The model for end-stage liver disease (MELD). *Hepatology* 2007;45:797–805.
- [6] Kim WR, Biggins SW, Kremers WK, et al. Hyponatremia and mortality among patients on the liver-transplant waiting list. *N Engl J Med* 2008;359:1018–1026.
- [7] Pugh RN, Murray-Lyon IM, Dawson JL, Pietroni MC, Williams R. Transection of the oesophagus for bleeding oesophageal varices. *Br J Surg* 1973;60:646–649.
- [8] Levesque E, Hoti E, Azoulay D, et al. Prospective evaluation of the prognostic scores for cirrhotic patients admitted to an intensive care unit. *J Hepatol* 2012;56:95–102.
- [9] Fine JP, Gray RJ. A proportional hazards model for the subdistribution of a competing risk. *J Am Stat Assoc* 1999;94:496–509.
- [10] Wolbers M, Koller MT, Wittman JC, Steyerberg EW. Prognostic models with competing risks. *Methods and application to coronary risk prediction. Epidemiology* 2009;20:555–561.
- [11] Harrell Jr FE, Lee KL, Mark DB. Multivariable prognostic models: issues in developing models, evaluating assumptions and adequacy, and measuring and reducing errors. *Stat Med* 1996;15:361–387.
- [12] Pencina MJ, D'Agostino RB. Overall C as a measure of discrimination in survival analysis: model specific population value and confidence interval estimation. *Stat Med* 2004;23:2109–2123.
- [13] Pencina MJ, D'Agostino Sr RB, D'Agostino Jr RB, Vasan RS. Evaluating the added predictive ability of a new marker: from area under the ROC curve to reclassification and beyond. *Stat Med* 2008;27:157–172.
- [14] Altman DG, Royston P. What do we mean by validating a prognostic model? *Stat Med* 2000;19:453–473.
- [15] Shawcross DL, Sharifi Y, Canavan JB, et al. Infection and systemic inflammation, not ammonia, are associated with Grade 3/4 hepatic encephalopathy, but not mortality in cirrhosis. *J Hepatol* 2011;54:640–649.
- [16] Jalan R, Stadlbauer V, Sen S, Cheshire L, Chang YM, Mookerjee RP. Role of predisposition, injury, response and organ failure in the prognosis of patients with acute-on-chronic liver failure: a prospective cohort study. *Crit Care* 2012;16:R227.
- [17] O'Grady JG, Alexander GJ, Hayllar KM, Williams R. Early indicators of prognosis in fulminant hepatic failure. *Gastroenterology* 1989;97:439–445.
- [18] Cazzaniga M, Dionigi E, Gobbo G, Fioretti A, Monti V, Salerno F. The systemic inflammatory response syndrome in cirrhotic patients: relationship with their in-hospital outcome. *J Hepatol* 2009;51:475–482.
- [19] Cholongitas E, Marelli L, Shusang V, et al. A systematic review of the performance of the model for end-stage liver disease (MELD) in the setting of liver transplantation. *Liver Transpl* 2006;12:1049–1061.