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Chapter 2

Genetic association analysis of LARS2 with type 2 diabetes

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Abstract

Aims/hypothesis

LARS2 has been previously identified as a potential type 2 diabetes susceptibility gene through the low frequency (LF) H324Q (rs71645922) variant (MAF 3.0%). However, this association did not achieve genome-wide levels of significance. The aim of this study was to establish the true contribution of this variant and common variants in LARS2 (MAF >5%) to type 2 diabetes risk.

Methods

We combined genome-wide association study (GWAS) data (n=10128) from the DIAGRAM consortium, with independent data derived from a tagging SNP approach in Dutch individuals (n=999), and took forward two SNPs of interest for replication in up to 11163 Dutch subjects. In addition, because inspection of the GWAS data identified a cluster of LF variants with evidence of type 2 diabetes association, we attempted replication of rs9825041 (a proxy for this group) and the previously identified H324Q variant in up to 35715 subjects of European descent. **Results**

No association of the common SNPs in *LARS2* with type 2 diabetes was found. Our replication studies for the 2 LF variants, rs9825041 and H324Q failed to confirm an association with type 2 diabetes in Dutch, Scandinavian and UK samples (OR 1.03 (0.95-1.12), p=0.45, n=31962 and OR 0.99 (0.90-1.08), p=0.78, n=35715 respectively).

Conclusion

In this study, the largest study examining the association of sequence variants in LARS2 with type 2 diabetes susceptibility we find no evidence to support previous data indicating a role in type 2 diabetes susceptibility.

Introduction

Changes in mitochondrial function are observed in patients with type 2 diabetes and their first degree relatives. Previous studies have indicated that genes involved in oxidative phosphorylation are down regulated in the muscle cells of type 2 diabetes patients (1). Furthermore, the muscle mitochondria from patients with type 2 diabetes have an impaired bioenergetic capacity (2). Mitochondria also play an important role in insulin secretion and sensitivity (3;4). Previously, our group has shown that a mutation in the mitochondrial DNA encoded tRNA-Leu(UUR) gene is associated with maternally inherited diabetes and deafness (5). In addition, an H324Q (rs71645922) variant in the nuclear encoded mitochondrial LARS2 gene has shown an association with type 2 diabetes in work previously carried out by our group (6). The LARS2 gene encodes for the mitochondrial leucyl tRNA synthetase (EC 6.1.1.4), which catalyzes the aminoacylation of both mitochondrial leucyl tRNAs with leucine and is therefore essential for mitochondrial protein synthesis. By analyzing the coding region for the LARS2 gene we found an H324Q (rs71645922) variant, and demonstrated an association with type 2 diabetes susceptibility in a meta-analysis of four independent cohorts from the Netherlands and Denmark (OR = 1.40 (95% CI 1.12-1.76, P = 0.004, n = 7836) (6). In recent years the advent of genome-wide association studies (GWAS) and the accumulation of large data sets capable of detecting associations to levels of genome wide significance appropriate for such studies ($p<5$ *10⁻⁸) has identified close to 20 loci impacting on type 2 diabetes susceptibility. However, low frequency variants such as H324Q are generally poorly captured by such studies. We set out therefore to re-evaluate the possible contribution of this LF variant to type 2 diabetes susceptibility in appropriately sized samples. We also used a combination of publicly available (DIAGRAM consortium) and newly derived tagging SNP data to undertake the most comprehensive assessment of the LARS2 locus yet performed.

Materials and methods

Study samples

The first part of our study was aimed at the identification of common alleles associated with increased type 2 diabetes susceptibility using DIAGRAM consortium data and a tagging SNP approach. For this we genotyped several European samples.

The first sample we included was from the Hoorn Study (here designated as NL1) (7). From this Dutch population based study from the city of Hoorn, in North-Western Netherlands, we selected 519 normal glucose tolerant (NGT) subjects and 480 type 2 diabetes subjects. Glucose tolerance was assessed using a fasting oral glucose tolerance test (OGTT), according to 1999 World Health Organization (WHO) criteria (8). This sample was used for the analysis of common variation in LARS2 with a tagging SNP approach. Variants in LARS2 identified from the DIAGRAM meta-analysis and the tagging SNP approach were then taken forward for replication in three Dutch samples, designated NL2, 3 and 4 respectively. The second sample from the Netherlands (designated NL2) included 1517 controls and 821 cases (9;10). The1517 controls were randomly selected from the New Hoorn Study (NHS), which is an ongoing, population based study from the city of Hoorn, which does not overlap with the original Hoorn Study (NL1). We included 147 cases from the NHS and the remainder of the cases (n=674) were recruited from the diabetes clinics of the Leiden University Medical Centre (LUMC, Leiden) and from the Vrije Universiteit medical centre (VUmc, Amsterdam). All subjects in this replication sample were Dutch Caucasians and all NGT subjects underwent an OGTT according to WHO criteria(8).

The third replication sample was ascertained from the Breda study (NL3) (11;12). This is a case control study from the city of Breda, in Southern Netherlands. The 920 controls were from the Dutch blood bank and self reported a non-diabetic state. The 501 cases had type 2 diabetes based on WHO criteria (8). For the fourth replication sample we selected 5183 NGT subjects and 1222 type 2 diabetes subjects from the population based ERGO study from the city of Rotterdam in the South-western region of the Netherlands (NL4) (13).

In total 8139 controls and 3024 type 2 diabetes cases were included in our replication study in the Netherlands.

The second part of this study was focused on the follow up of 2 low frequency variants in LARS2 and for this we carried out replication in samples from the Netherlands (NL1-4) as well as samples from the UK (UK1,2), Denmark (DK1), Finland (FI1,2) and Sweden (SE1).

We included one replication sample from Denmark (designated DK1) (14). This sample consists of 514 NGT controls which are randomly selected from public registers at the Steno Diabetes Center and the Research Centre for Prevention and Health, Copenhagen, Denmark. The 706 cases were recruited from the Steno Diabetes Center. NGT subjects underwent an OGTT according to WHO criteria (8). Two UK samples were included. The first (UK1) was the UKT2DGC (United Kingdom Type 2 Diabetes Genetics Consortium) case-control sample comprising 4124 type 2 diabetes cases and 5126 controls ascertained in Tayside, Scotland. Details of the ascertainment scheme and recruitment criteria for this sample have been described elsewhere (15;16): the enlarged sample used here represents continuing recruitment to this resource under precisely the same criteria. The second sample (UK2) consists of 1853 type 2 diabetes cases ascertained as part of the BDA Warren 2 collection (Exeter, London, Oxford, Norwich and Newcastle) and 10220 control samples. The latter represent the full British 1958 Birth Cohort (n=7133) and The United Kingdom Blood Services Collection of Common Controls (UKBS) (n= 3087), a subset of which featured in the WTCCC genome wide association scan (both samples were collected throughout the UK) (15;16). Finally, we included samples from Finland and Sweden. One was a case-control sample from the Botnia region of Finland, here designated as FI1. This sample consisted of 353 controls and 402 cases. The second sample originated from Sweden (Skara and Malmö), here designated as the Swedish case-control study (SE1) and consisted of 468 controls and 480 cases (17;18). Furthermore, we included a set of trios originating from the Botnia region of Finland (FI2). This sample consisted of 211 probands (multiple diabetic sibs) and 370 parents. All study samples are summarized in table 1.

In total 25191 controls and 10800 type 2 diabetes cases were included for the follow up of the LF variants.

All studies were approved by the appropriate medical ethical committees and were in accordance with the principles of the Declaration of Helsinki. All participants provided written, informed consent for this study.

Common SNP selection

Common SNPs (MAF >5%) in the LARS2 locus were selected for follow-up based on the data from the DIAGRAM meta-analysis (gene boundaries chr3: 45373001…45698001) (22). SNPs with a P<0.05 were genotyped in the Dutch replication samples (NL1-4). Furthermore, tagging SNPs in LARS2 were selected for genotyping in the NL1 sample using the HapMap database and Tagger software (19;20) (selection criteria and SNPs shown in supplementary table S1).

Table 1. Description of study samples.

n.a. not available

a. Based on the British 1958 Birth Cohort (7133) and Panel 2 of the United Kingdom Blood Services Collection of Common Controls (n=1643)

b. parents

c. probands

Genotyping and quality control

SNPs selected for follow-up in our replication samples were genotyped using Taqman SNP genotyping assays (Applied Biosystems, Foster City, USA). Tagging SNPs were genotyped in the NL1 sample using the Sequenom platform (Sequenom, San Diego, USA). Assays showing overlapping clusters, success rates below 95% or not obeying Hardy Weinberg Equilibrium (HWE) (p<0.05) were excluded from analysis. Duplicate samples (~5%) showed complete concordance.

Statistical analysis

Differences in genotype distribution and allele frequencies were analyzed using a chi-squared test. ORs were calculated using an additive model, which was the best fit for the data. Homogeneity of ORs between the different samples was calculated with a Tarone's test after which a common OR was calculated with a Mantel-Haenszel test using a fixed effects model. Results from OGTT (only normal glucose tolerant subjects) were analyzed with univariate analysis of variance, using additive, recessive and dominant models and correction for age, BMI and gender as possible confounders. Association in the Botnia trios was assessed by the transmission disequilibrium test (TDT). All general statistics were calculated using SPSS 16.0 (SPSS Inc, Chicago, USA). For statistics involving the geographical distribution of the H324Q (rs71645922) variant in the UK population (described below) we used StatXact v 6.0 (Cytel software corps, Cambridge, MA, USA). Power calculations were performed using Quanto (21). From the DIAGRAM consortium meta-analysis of common variants we selected for replication all common SNPs with a p<0.05. At this alpha the DIAGRAM consortium metaanalysis had at least 80% power to detect a variant with OR \geq 1.20 (MAF >0.05) (22). Combined with our Dutch replication sample we had at least 80% power to replicate the association of a variant with an $OR \ge 1.09$ at the observed MAFs of 0.19 (rs952621) and 0.24 (rs17637703) respectively (alpha = 0.05) or OR ≥1.12 at alpha = 10^{-4}). Power of the tagging SNP approach in NL1 was limited (80% power to detect a variant with an OR 21.6 (alpha=0.05, MAF=0.05) or OR \geq 1.45 at the observed lowest MAF of 0.10) therefore we replicated in NL2-4 only our strongest signal from the NL1 sample ($rs17637703$, $p = 0.07$).

Whilst extensive GWAS have indicated that the effect sizes of common variants influencing type 2 diabetes risk are modest, the potential remains for low-frequency variants to have effects on type 2 diabetes risk that are more substantial which was corroborated by our previous observation regarding the H324Q variant (6). Power calculations at the start of the project demonstrated that we had at least 99% power to detect an effect size similar to our initial finding concerning H324Q (rs71645922) (OR 1.4) and at least 80% power to detect an OR of 1.13 (α =0.05) (6). From the DIAGRAM meta-analysis we used an alpha of 0.05 to select other LF SNPs for replication. At this alpha the power in DIAGRAM was 80% to detect association for variants with ORs ranging from 1.24 (MAF = 0.03) to 1.45 (MAF = 0.01 and alpha $= 0.05$). For replication of the two low frequency variants (observed MAFs ~0.03(H324Q, rs71645922) and ~0.05 (rs9825041) respectively) we had in our complete replication sample at least 80% power to detect an OR \geq 1.13 (25191 controls and 10800 type 2 diabetes cases and alpha $= 0.05$).

Results

Common LARS2 variants in available DIAGRAM GWAS data

We analyzed the data from the DIAGRAM GWAS meta-analysis (22) for the LARS2 gene (100% coverage (MAF $>$ 5%), according to HapMap phase 2, April 2007, CEU population) and observed 1 common SNP (rs952621, directly typed) showing weak evidence of association with type 2 diabetes ($OR = 1.11$ (1.02 – 1.20), $p = 0.01$ for the T allele). This SNP was also captured in our complementary tagging SNP approach (NL1) and we found an OR of 1.13 (0.89 – 1.43), $p = 0.33$ for the same allele. However, additional genotyping in the Dutch samples (NL2,4) and meta-analysis of all data resulted in a common OR of 1.05 (0.99 – 1.11), $p =$ 0.13 (n=19870). As there was no convincing evidence of association in our samples, this SNP was not analysed further.

No other common SNP in LARS2 showed evidence for association with type 2 diabetes in the GWAS data. The same was true of the tagging SNP analysis conducted in the NL1 sample (supplementary table s1). In the latter analysis, rs17637703 showed weak evidence of association (OR = 1.22 (95% C.I. 0.99 – 1.50), $p = 0.07$ but this was not confirmed in the Dutch replication samples (the common OR was 0.98 ($0.91 - 1.06$), $p = 0.62$ (n = 10087), in line with the DIAGRAM result for this SNP (OR 1.02 (0.94 – 1.10)).

Low frequency variants in LARS2

In addition to the common variants, the DIAGRAM meta-analysis also captured fourteen LF SNPs (0.01<MAF<0.05) within the LARS2 gene, ten of which are in high LD with each other (r^2 > 0.95 according to HapMap, supplementary figure S1) and showed some evidence for association with type 2 diabetes (ORs 1.17 – 1.21; p $0.02 - 0.05$). We selected rs9825041 (OR = 1.20 $(1.03 - 1.39)$, p = 0.02) as a proxy for the group for genotyping in the replication samples but no association with type 2 diabetes was observed (table 2). Homogeneity of ORs was tested with a Tarone's test ($p = 0.67$) and we calculated a common OR across all studies of 1.03 $(0.95 - 1.12)$, $p = 0.45$, $(8959 \text{ cases}, 23003 \text{ controls})$.

Follow up of the H324Q (rs71645922) variant in LARS2

Finally, we examined the association of the H324Q (rs71645922) variant with type 2 diabetes in our replication samples from the Netherlands (NL2-4), UK (UK1, 2), Sweden (SE1) and Finland (FI1). This variant was not captured by the GWAS and was not captured by any of the SNPs mentioned above (r^2 < 0.17). Subjects in the Dutch replication samples that were included in our original study of this variant were excluded from analysis ($n = 914$ from the NL4 study). The replication samples did not confirm our previously observed association. A meta-analysis of all available studies including our previous data from the Netherlands (NL1) and Denmark (DK1, 2) (6) resulted in an overall OR of 0.99 (0.90 – 1.08), $p = 0.78$, $n =$ 35715 subjects (10399 type 2 diabetes subjects, table 2). In addition, we did not observe a significant excess of transmission of the risk allele in the Botnia trios (FI2, transmitted / untransmitted = $18/14$, OR 1.29 (0.64 – 2.59), p = 0.48). To investigate possible heterogeneity between the studies we performed several analyses. For age stratification we created, based on the age distribution in the Dutch samples the following age strata; ≤ 60 years, 61-70 years and >70 years. A decreased frequency of the risk allele was observed in type 2 diabetes subjects with increasing age in most but not all samples (data not shown, available on

request), but this did not reach statistical significance. Furthermore we looked at age at diagnosis of type 2 diabetes and allele frequencies in those with early onset diabetes (\leq 45 years) and those with an age at diagnosis above 45 years. Although the allele frequency was slightly higher in those with early onset diabetes this was not statistically significant nor was the age at diagnosis in carriers and non-carriers (all $p > 0.05$, data not shown). Stratification for gender and BMI (where data available) did not affect the outcome in the Dutch studies (NL1-NL4, data not shown) and was therefore not further investigated.

H324Q (rs71645922) shows marked variation in MAF between the various European-descent samples examined (control MAF ranges from 1.9% to 4.8%). In the two large UK control samples, for example, there was a highly significant ($p = 5$) 10^{-7} , using an exact implementation of the Cochran Armitage trend test) difference in allele frequencies between UK1 (recruited exclusively in Scotland) and UK2 (recruited throughout the UK) that made us consider the possibility that this variant was showing variation in allele frequency along the south-north cline previously described in the WTCCC study and others (16;23-26).To test this, we made use of information on the region of ascertainment available for the UK 1958 Birth Cohort and UK Blood Service and analysed genotype frequencies based on subdivisions of the UK into 4 major regions, namely (1) Scotland, (2) Northern England(Yorkshire and the Humber, North East, North West), (3) UK Midlands (East Midlands, West Midlands, Wales, East of England) and (4) South of England (South East, Greater London, and South West). We found some evidence (supplementary figure S2) for a North - South gradient across the UK. (MAF 4.66 %, 3.41%, 3.31%, and 3.28% respectively) with $p = 0.038$ calculated using the Jonckheere-Terpstra Test (StatXact v 6.0; Cytel software corporation, Cambridge, MA, USA). No such MAF gradient was observed in other European samples (supplementary figure S3)

Minor allele frequencies are shown for cases and controls.

n.m.: not measured

M-A: meta-analysis

a. Data taken from reference (6). Results from the NL4 study are partially from this previous research ($n =$ 914).

b. Meta analysis of the DIAGRAM consortium GWAS, $n = 10128$ (4549 cases and 5579 controls) (22).

Discussion

We found no evidence of common SNPs in *LARS2* being associated with type 2 diabetes in our samples. We therefore conclude that it is unlikely that common SNPs in LARS2 are associated with type 2 diabetes susceptibility.

Several low frequency SNPs, which are all in high LD with each other $(r^2 > 0.95)$ showed nominal evidence of association with type 2 diabetes in the DIAGRAM meta-analysis. However, we have been unable to confirm this association in our large replication samples from the Netherlands, Denmark and the UK ($n = 31962$) and therefore conclude that the nominal p-values observed in the GWAS are most likely consistent with statistical noise.

The previously observed association of the H324Q (rs71645922) variant in LARS2 with type 2 diabetes was not confirmed in our replication samples (table 2). Power

LARS2 and type 2 diabetes

calculations at the start of the project demonstrated that we had at least 99% power to detect an effect size similar to our initial finding concerning H324Q (rs71645922) (OR 1.4) and at least 80% power to detect an OR of 1.13. Before excluding the previous association as false, we considered the possibility of heterogeneity but found no evidence that age, age at diagnosis, BMI and gender were responsible. Another possibility, raised by the evidence for variation in H324Q MAF across the UK is that the previous association in Dutch and Danish subjects reflected the effects of hidden population structure. It seems, however, unlikely that population stratification effects were responsible for the original reports of H324Q (rs71645922) associations as the cases and controls in that study were recruited from the same relatively narrow geographic regions within the Netherlands and Denmark (6). Also in our additional Dutch replication cohorts we could not detect a MAF gradient across the country. However, there are differences in MAF between different countries (supplementary Figure S2, S3). Migration patterns in the UK appear to reflect an increase in the MAF of H324Q (rs71645922) and therefore this may be a potential migration marker, however this needs to be demonstrated in other populations. Since stratification for BMI and gender did not affect our result, we can exclude that these variables confounded our observation. The reason for the discrepancy between our first and current study is likely to reflect chance. Three other low frequency non-synonymous SNPs are present in the LARS2 locus; K727N (rs36054230), E831D (rs9827689) and E868K (rs34965084). However, according to dbSNP and our own sequencing efforts (6) these SNPs are only identified in the African population and not polymorphic in the European population. Therefore, these additional non-synonymous variants were not analysed in this study. As our study does not include a thorough resequencing of the complete LARS2 locus we cannot fully exclude that other yet unknown LF variants are present and associated with type 2 diabetes. Results from the 1000 genomes project should facilitate a thorough investigation of LF SNPs in LARS2 in the future. In conclusion, our findings do not support the hypothesis that common variants in LARS2 are major type 2 diabetes susceptibility factors. We have also conducted one of the largest (up to 35715 subjects) replication studies for two low frequency variants for type 2 diabetes and also these data do not support a significant role as

type 2 diabetes susceptibility variants. We therefore conclude that currently known genetic variation in LARS2 does not play an important role in type 2 diabetes susceptibility.

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Supplementary figure S1. LD plot of all known LF SNPs in the LARS2 locus.

LD plot of all known LF SNPs in LARS2 (MAF \leq 0.05; HAPMAP phase 2, CEU). r^2 for each of the SNP pairs is given (black r^2 =1, shades of gray 0< r^2 <1 and white r^2 =0).

The 14 LF SNPs directly genotyped in DIAGRAM are boxed. In the insert only the 14 LF SNPs included in the DIAGRAM scan are shown.

rs9825041 was used as a proxy for the block and is shown in the box.

Supplementary figure S2. Geographic stratification in the UK for the H324Q (rs71645922) variant.

The numbers represent minor allele frequencies of the H324Q (rs71645922) variant in each of the geographic regions.

Supplementary figure S3. Geographic stratification in other European samples.

The numbers is the circles represent minor allele frequencies of H324Q (rs71645922) in each of the study samples.

Supplementary table S1. Overview of tagging SNPs, including results from the NL1 study.

The LARS2 locus was also investigated using a tagging SNP approach in the NL1 sample. Common SNPs (MAF > 0.05), in LARS2 were identified using HAPMAP (gene boundaries set at 45,405,079 - 45,565,332)[19].

Tagging SNPs were selected using the Tagger option in Haploview (NCBI build 36, phase II, April 2007, population CEU, LOD threshold 3.0 and $r^2 > 0.8$ [20].

Allele 1 is the major and allele 2 is the minor allele.

n.a.: not available.

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