

CENTRAL LIMIT THEOREM FOR THE EDWARDS MODEL

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The Edwards model in one dimension is a transformed path measure for standard Brownian motion discouraging self-intersections. We prove a central limit theorem for the endpoint of the path, extending a law of large numbers proved by Westwater. The scaled variance is characterized in terms of the largest eigenvalue of a one-parameter family of differential operators, introduced and analyzed by van der Hofstad and den Hollander. Interestingly, the scaled variance turns out to be independent of the strength of self-repulsion and to be strictly smaller than one (the value for free Brownian motion).

0. Introduction and main result.

0.1. *The Edwards model.* Let $(B_t)_{t \geq 0}$ be standard one-dimensional Brownian motion starting at 0. Let P denote its distribution on path space and E the corresponding expectation. The *Edwards model* is a transformed path measure discouraging self-intersections, defined by the intuitive formula

$$(0.1) \quad \frac{dP_T^\beta}{dP} = \frac{1}{Z_T^\beta} \exp \left[-\beta \int_0^T ds \int_0^T dt \delta(B_s - B_t) \right], \quad T \geq 0.$$

Here δ denotes Dirac's function, $\beta \in (0, \infty)$ is the *strength of self-repulsion* and Z_T^β is the normalizing constant.

A rigorous definition of P_T^β is given in terms of Brownian local times as follows. It is well known [see Revuz and Yor (1991), Section VI.1] that there exists a jointly continuous version of the Brownian local time process $(L(t, x))_{t \geq 0, x \in \mathbb{R}}$ satisfying the occupation time formula

$$(0.2) \quad \int_0^t f(B_s) ds = \int_{\mathbb{R}} L(t, x) f(x) dx \quad P\text{-a.s.} \quad (f: \mathbb{R} \rightarrow \mathbb{R}^+ \text{ Borel, } t \geq 0).$$

Think of $L(t, x)$ as the amount of time the Brownian motion spends in x until time t . The Edwards measure in (0.1) may now be defined by

$$(0.3) \quad \frac{dP_T^\beta}{dP} = \frac{1}{Z_T^\beta} \exp \left[-\beta \int_{\mathbb{R}} L(T, x)^2 dx \right],$$

where $Z_T^\beta = E(\exp[-\beta \int_{\mathbb{R}} L(T, x)^2 dx])$ is the normalizing constant. The random variable $\int_{\mathbb{R}} L(T, x)^2 dx$ is called the *self-intersection local time*. Think of

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this as the amount of time the Brownian motion spends in self-intersection points until time T .

The path measure P_T^β is the continuous analogue of the self-repellent random walk (called the *Domb–Joyce model*), which is a transformed path measure for the discrete simple random walk. The latter is used to study the long-time behavior of random polymer chains. The effect of the self-repulsion is of particular interest. This effect is known to spread out the path on a linear scale (i.e., B_T is of order T under the law P_T^β as $T \rightarrow \infty$). It is the aim of this paper to study the fluctuations of B_T around the linear asymptotics. Our main result appears in Theorem 2.

0.2. *Theorems.* The starting point of our paper is the following law of large numbers.

THEOREM 1 [Westwater (1984)]. *For every $\beta \in (0, \infty)$ there exists a $\theta^*(\beta) \in (0, \infty)$ such that*

$$(0.4) \quad \lim_{T \rightarrow \infty} P_T^\beta \left(\left| \frac{B_T}{T} - \theta^*(\beta) \right| \leq \varepsilon \mid B_T > 0 \right) = 1 \quad \text{for every } \varepsilon > 0.$$

[By symmetry, (0.4) says that the distribution of B_T/T under P_T^β converges weakly to $\frac{1}{2}(\delta_{\theta^*(\beta)} + \delta_{-\theta^*(\beta)})$ as $T \rightarrow \infty$, where δ_θ denotes the Dirac point measure at $\theta \in \mathbb{R}$.]

Theorem 1 says that the self-repulsion causes the path to have a ballistic behavior no matter how weak the interaction. Westwater (1984) proved this result by applying the Ray–Knight representation for Brownian local times and using large deviation arguments.

The speed $\theta^*(\beta)$ was characterized by Westwater in terms of the smallest eigenvalue of a certain differential operator. In the present paper, however, we prefer to work with a different operator, introduced and analyzed in van der Hofstad and den Hollander (1995). For $a \in \mathbb{R}$, define $\mathcal{K}^a: L^2(\mathbb{R}_0^+) \cap C^2(\mathbb{R}_0^+) \rightarrow C(\mathbb{R}_0^+)$ by

$$(0.5) \quad (\mathcal{K}^a x)(u) = 2ux''(u) + 2x'(u) + (au - u^2)x(u)$$

for $u \in \mathbb{R}_0^+ = [0, \infty)$. The Sturm–Liouville operator \mathcal{K}^a will play a key role in the present paper. It is symmetric and has a largest eigenvalue $\rho(a)$ with multiplicity 1. The map $a \mapsto \rho(a)$ is real-analytic, strictly convex and strictly increasing, with $\rho(0) < 0$, $\lim_{a \rightarrow -\infty} \rho(a) = -\infty$ and $\lim_{a \rightarrow \infty} \rho(a) = \infty$. [The operator \mathcal{K}^a is a scaled version of the operator \mathcal{L}^a originally analyzed in van der Hofstad and den Hollander (1995), Section 5, namely $(\mathcal{K}^a x)(u) = (\mathcal{L}^a \bar{x})(u/2)$ where $\bar{x}(u) = x(2u)$.]

Define a^* , b^* , $c^* \in (0, \infty)$ by

$$(0.6) \quad \rho(a^*) = 0, \quad b^* = \frac{1}{\rho'(a^*)}, \quad c^{*2} = \frac{\rho''(a^*)}{\rho'(a^*)^3}.$$

Our main result is the following central limit theorem.

THEOREM 2. For every $\beta \in (0, \infty)$ there exists a $\sigma^*(\beta) \in (0, \infty)$ such that

$$(0.7) \quad \lim_{T \rightarrow \infty} P_T^\beta \left(\frac{B_T - \theta^*(\beta)T}{\sigma^*(\beta)\sqrt{T}} \leq C \mid B_T > 0 \right) = \mathcal{N}((-\infty, C]) \quad \text{for all } C \in \mathbb{R},$$

where \mathcal{N} denotes the normal distribution with mean 0 and variance 1. The scaled mean and variance are given by

$$(0.8) \quad \theta^*(\beta) = b^* \beta^{1/3}, \quad \sigma^*(\beta) = c^*.$$

Theorem 2 says that the fluctuations around the asymptotic mean have the classical order \sqrt{T} , are symmetric, and even do not depend on the interaction strength.

The numerical values of the constants in (0.6) are

$$(0.9) \quad a^* = 2.189 \pm 0.001, \quad b^* = 1.11 \pm 0.01, \quad c^* = 0.7 \pm 0.1.$$

The values for a^* and b^* were obtained in van der Hofstad and den Hollander (1995), Section 0.5, by estimating $\rho(a)$ for a range of a -values. This can be done very accurately via a discretization procedure. (A rigorous upper bound for a^* is given in Lemma 6 in Section 4.1.) The same data produce the value for c^* . Note that $c^* < 1$. Apparently, as the path is pushed out to infinity, its fluctuations are squeezed compared to those of the free motion with $\theta^*(0) = 0$, $\sigma^*(0) = 1$.

0.3. Scaling in β . It is noteworthy that the scaled mean depends on β in such a simple manner and that the scaled variance does not depend on β at all. These facts are direct consequences of the Brownian scaling property. Namely, we shall deduce from (0.7) that for every $\beta \in (0, \infty)$,

$$(0.10) \quad \theta^*(\beta) = \theta^*(1)\beta^{1/3}, \quad \sigma^*(\beta) = \sigma^*(1).$$

Indeed, for $a, T > 0$,

$$(0.11) \quad (B_T, (L(T, x))_{x \in \mathbb{R}}) =_{\mathcal{G}} (a^{-1/2}B_{aT}, (a^{-1/2}L(aT, a^{1/2}x))_{x \in \mathbb{R}}),$$

where $=_{\mathcal{G}}$ means equality in distribution [see Revuz and Yor (1991), Chapter VI, Example (2.11), 1°]. Apply this to $a = \beta^{2/3}$ to obtain, via (0.3), that

$$(0.12) \quad P_T^\beta (B_T)^{-1} = P_{\beta^{2/3}T}^1 (\beta^{-1/3}B_{\beta^{2/3}T})^{-1},$$

where we write $\mu(X)^{-1}$ for the distribution of a random variable X under a measure μ . In particular, we have for all $C \in \mathbb{R}$,

$$(0.13) \quad \begin{aligned} P_T^\beta \left(\frac{B_T - \theta^*(1)\beta^{1/3}T}{\sigma^*(1)\sqrt{T}} \leq C \mid B_T > 0 \right) \\ = P_{\beta^{2/3}T}^1 \left(\frac{B_{\beta^{2/3}T} - \theta^*(1)\beta^{2/3}T}{\sigma^*(1)\sqrt{\beta^{2/3}T}} \leq C \mid B_{\beta^{2/3}T} > 0 \right). \end{aligned}$$

The r.h.s. tends to $\mathcal{N}((-\infty, C])$ as $T \rightarrow \infty$ [in (0.7) pick $\beta = 1$ and replace T by $\beta^{2/3}T$]. Since the pair $(\theta^*(\beta), \sigma^*(\beta))$ is uniquely determined by (0.7), we arrive at (0.10).

0.4. *Outline of the proof.* Theorem 2 is the continuous analogue of the central limit theorem for the Domb–Joyce model proved by König (1996). We shall be able to use the skeleton of that paper, but the Brownian context will require new ideas and methods. The remaining sections are devoted to the proof of Theorem 2. We give a short outline.

In Section 1, we use the well-known Ray–Knight theorems for the local times of Brownian motion to express the l.h.s. of (0.7) in terms of two- and zero-dimensional squared Bessel processes. The former describes the local times in the area $[0, B_T]$; the latter describes the local times in $(-\infty, 0]$ (respectively, $[B_T, \infty)$).

In Section 2, with the help of some analytical properties of the operator \mathcal{K}^a proved in van der Hofstad and den Hollander (1995), we introduce a Girsanov transformation of the two-dimensional squared Bessel process. The goal of this transformation is to absorb the random variable $\exp(-\beta \int_0^{B_T} L(T, x)^2 dx)$ into the transition probabilities. The transformed process turns out to have strong recurrence properties. The Gaussian behavior of $(B_T - \theta^*(\beta)T)/\sqrt{T}$ is traced back to the asymptotic normality of the *inverse* of a certain additive functional of this transformed process. Thus, the central limit behavior is determined by those parts of the Brownian path that fall in the area $[0, B_T]$.

In Section 3, we prove a central limit theorem for the inverse process. Furthermore, as a second important ingredient in the proof, we derive a limit law and a rate of convergence result for the composition of the transformed process with the inverse process.

In Section 4, we finish the proof of Theorem 2 by showing that the contribution of the local times in $(-\infty, 0] \cup [B_T, \infty)$ remains bounded as $T \rightarrow \infty$ and is therefore cancelled by the normalization in the definition of the transformed path measure in (0.3).

1. Brownian local times. Since the dependence on β has already been isolated [see (0.13)], we may and shall restrict to the case $\beta = 1$.

Throughout the sequel we shall frequently refer to Revuz and Yor (1991), Karatzas and Shreve (1991), van der Hofstad and den Hollander (1995). We shall therefore adopt the abbreviations RY, KS and HH for these references.

The remainder of this paper is devoted to the proof of the following key proposition.

PROPOSITION 1. *There exists an $S \in (0, \infty)$ such that for all $C \in \overline{\mathbb{R}}$,*

$$(1.1) \quad \lim_{T \rightarrow \infty} \exp(a^*T) E \left(\exp \left(- \int_{\mathbb{R}} L(T, x)^2 dx \right) 1_{0 < B_T \leq b^*T + C\sqrt{T}} \right) = S \mathcal{N}_{c^*2}((-\infty, C]),$$

where a^* , b^* and c^* are defined in (0.6), and \mathcal{N}_{σ^2} denotes the normal distribution with mean 0 and variance σ^2 .

Theorem 2 follows from Proposition 1, since it implies that the conditional distribution of $(B_T - b^*T)/\sqrt{T}$ given $B_T > 0$ converges to $\mathcal{N}_{c^{*2}}$ [divide the l.h.s. of (1.1) by the same expression with $C = \infty$ and recall (0.3)].

Sections 1.1 and 1.2 contain preparatory material. Section 1.3 contains the key representation in terms of squared Bessel processes on which the proof of Proposition 1 will be based.

1.1. Ray-Knight theorems. This subsection contains a description of the *time-changed* local time process in terms of squared Bessel processes. The material being fairly standard, our main purpose is to introduce appropriate notation and to prepare for Lemma 1 in Section 1.2 and Lemma 2 in Section 1.3.

For $u \in \mathbb{R}$ and $h \geq 0$, let τ_h^u denote the time change associated with $L(t, u)$; that is,

$$(1.2) \quad \tau_h^u = \inf\{t > 0: L(t, u) > h\}.$$

Obviously, the map $h \mapsto \tau_h^u$ is right-continuous and increasing, and therefore makes at most countably many jumps for each $u \in \mathbb{R}$. Moreover, $P(L(\tau_h^u, u) = h \text{ for all } u \geq 0) = 1$ (see RY, Chapter VI). The following lemma contains the well-known Ray-Knight theorems. It identifies the distribution of the local times at the random time τ_h^u as a process in the spatial variable running forwards, respectively backwards, from u . We write $C_c^2(\mathbb{R}^+)$ to denote the set of twice continuously differentiable functions on $\mathbb{R}^+ = (0, \infty)$ with compact support.

RK THEOREMS. Fix $u, h \geq 0$. The random processes $(L(\tau_h^u, u + v))_{v \geq 0}$ and $(L(\tau_h^u, u - v))_{v \geq 0}$ are independent Markov processes, both starting at h .

(i) $(L(\tau_h^u, u + v))_{v \geq 0}$ is a zero-dimensional squared Bessel process ($BESQ^0$) with generator

$$(1.3) \quad (G^*f)(v) = 2vf''(v), \quad f \in C_c^2(\mathbb{R}^+).$$

(ii) $(L(\tau_h^u, u - v))_{v \in [0, u]}$ is the restriction to the interval $[0, u]$ of a two-dimensional squared Bessel process ($BESQ^2$) with generator

$$(1.4) \quad (Gf)(v) = 2vf''(v) + 2f'(v), \quad f \in C_c^2(\mathbb{R}^+).$$

(iii) $(L(\tau_h^u, -v))_{v \geq 0}$ has the same transition probabilities as the process in (i).

For the proof, see RY, Sections XI.1-2 and KS, Sections 6.3 and 6.4.

1.2. *The distribution of $((L(T, x))_{x \in \mathbb{R}}, B_T)$.* The RK theorems give us a nice description of the local time process at certain stopping times. In order to apply them to (0.3), we need to go back to the fixed time T . This causes some complications (e.g., we must handle the global restriction $\int_{\mathbb{R}} L(T, x) dx = T$), but these may be overcome by an appropriate conditioning.

This subsection contains a formal description of the joint distribution of the three random processes

$$(1.5) \quad (L(T, B_T + x))_{x \geq 0}, \quad (L(T, B_T - x))_{x \in [0, B_T]}, \quad (L(T, -x))_{x \geq 0},$$

in terms of the squared Bessel processes. The main intuitive idea is that, up to a P -null set [recall (1.2)],

$$(1.6) \quad \{\tau_h^u = T\} = \{B_T = u, L(T, B_T) = h\} \quad \text{for all } u, h \geq 0.$$

This has two consequences.

1. Conditioned on $\{B_T = u, L(T, B_T) = h\}$, the three processes in (1.5) are the squared Bessel processes from the RK theorems conditioned on having total integral equal to T .
2. The distribution of $(B_T, L(T, B_T))$ can be expressed in terms of the squared Bessel processes.

We shall make this precise in Lemma 1 below.

Before we proceed, let us briefly mention some earlier works on the distribution of $(L(T, x))_{x \in \mathbb{R}}$ with $T \geq 0$ either fixed or random independent of the motion. Perkins (1982) proves that $(L(1, x))_{x \in \mathbb{R}}$ is a semimartingale. Jeulin (1985) uses stochastic calculus, in particular Tanaka's formula, to recover the RK theorems and Perkins' result and to prove the conditional Markov property in x of the triple $(L(1, x), x \wedge B_1, \int_{-\infty}^x L(1, u) du)$ given $\inf_{s \leq 1} B_s$. In Biane and Yor (1988) the RK theorems are extended to the case where T is an exponentially distributed random time, independent of the Brownian motion, under the conditional law $P(\cdot | L(T, 0) = s, B_T = a)$ for any fixed $s, a > 0$. Finally, Biane, Le Gall and Yor (1987) also deal with the intuitive idea (1.6) when identifying the law of the process $((1/\sqrt{\tau_h^0})B_{u\tau_h^0})_{u \in [0, 1]}$.

Let us now return to our identification of the law of the process $((L(T, x))_{x \in \mathbb{R}}, B_T)$. In order to formulate the details, we must first introduce some notation. For the remainder of this paper, let

$$(1.7) \quad (X_v)_{v \geq 0} = \text{BESQ}^2, \quad (X_v^*)_{v \geq 0} = \text{BESQ}^0.$$

Note that $(X_v)_{v \geq 0}$ is recurrent and has 0 as an entrance boundary, while $(X_v^*)_{v \geq 0}$ is transient and has 0 as an absorbing boundary (see RY, Section XI.1). Denote by \mathbb{P}_h and \mathbb{P}_h^* the distributions of the respective processes conditioned on starting at $h \geq 0$. Denote the corresponding expectations by \mathbb{E}_h , respectively \mathbb{E}_h^* . Furthermore, define the following additive functional of BESQ^2 and its

time change:

$$(1.8) \quad \begin{aligned} A(u) &= \int_0^u X_v \, dv, \quad u \geq 0, \\ A^{-1}(t) &= \inf\{u > 0: A(u) > t\}, \quad t \geq 0. \end{aligned}$$

Note that both $u \mapsto A(u)$ and $t \mapsto A^{-1}(t)$ are continuous and strictly increasing towards infinity \mathbb{P}_h -a.s. So A and A^{-1} are in fact inverse functions of each other. We also need the analogous functional for BESQ^0 :

$$(1.9) \quad \begin{aligned} A^*(u) &= \int_0^u X_v^* \, dv, \quad u \in [0, \infty], \\ A^{*-1}(t) &= \inf\{u \geq 0: A^*(u) > t\}, \quad t \geq 0. \end{aligned}$$

Note that, \mathbb{P}_h^* -a.s., $u \mapsto A^*(u)$ is strictly increasing on the time interval $[0, \xi_0]$, where $\xi_0 = \inf\{v \geq 0: X_v^* = 0\} < \infty$ denotes the absorption time of BESQ^0 . Define Lebesgue densities φ_h and $\psi_{h_1, t}$ by

$$(1.10) \quad \begin{aligned} \varphi_h(t) \, dt &= \mathbb{P}_h^*(A^*(\infty) \in dt), \\ \psi_{h_1, t}(u, h_2) \, du \, dh_2 &= \mathbb{P}_{h_1}(A^{-1}(t) \in du, X_u \in dh_2) \end{aligned}$$

for a.e. $h, t, h_1, u, h_2 \geq 0$. (The function φ_h is explicitly identified in Lemma 7 in Section 4.2.) Put the quantities defined in (1.8)–(1.10) equal to zero if any of the variables are negative. Now the joint distribution of the three processes in (1.5) can be described as follows.

LEMMA 1. *Fix $T > 0$. For all nonnegative Borel functions Φ_1, Φ_2 and Φ_3 on $C(\mathbb{R}_0^+)$ and for any interval $I \subset [0, \infty)$,*

$$(1.11) \quad \begin{aligned} &E\left(\Phi_1((L(T, B_T + x))_{x \geq 0}) \Phi_2((L(T, -x))_{x \geq 0}) \right. \\ &\quad \left. \times \Phi_3((L(T, B_T - x))_{x \in [0, B_T]}) 1_{B_T \in I}\right) \\ &= \int_I du \int_{[0, \infty)^4} dt_1 \, dh_1 \, dt_2 \, dh_2 \\ &\quad \times \prod_{i=1}^2 \mathbb{E}_{h_i}^*(\Phi_i((X_v^*)_{v \geq 0}) \mid A^*(\infty) = t_i) \varphi_{h_i}(t_i) \\ &\quad \times \mathbb{E}_{h_1}(\Phi_3((X_v)_{v \in [0, u]}) \mid A^{-1}(T - t_1 - t_2) = u, X_u = h_2) \\ &\quad \times \psi_{h_1, T-t_1-t_2}(u, h_2). \end{aligned}$$

PROOF. Essentially, Lemma 1 is a formal rewrite using (1.8), (1.10) and the RK theorems, which say that under \mathbb{P}_h , respectively \mathbb{P}_h^*

$$(1.12) \quad \begin{aligned} (X_v)_{v \in [0, u]} &=_{\mathcal{D}} (L(\tau_h^u, u - v))_{v \in [0, u]} \\ (X_v^*)_{v \geq 0} &=_{\mathcal{D}} (L(\tau_h^u, u + v))_{v \geq 0}. \end{aligned}$$

However, the details are far from trivial.

We proceed in four steps, the first of which makes (1.6) precise and is the most technical.

STEP 1. $P(\tau_h^u \in dT) du dh = P(B_T \in du, L(T, B_T) \in dh) dT$ for a.e. $u, h, T \geq 0$.

PROOF. From the occupation time formula (0.2) we have for every $t \geq 0$,

$$(1.13) \quad \int_0^t 1_{B_s \in du} ds = L(t, u) du.$$

Hence, we obtain for all bounded and measurable functions $f: (\mathbb{R}^+)^2 \rightarrow \mathbb{R}$ and $g: \mathbb{R}^+ \rightarrow \mathbb{R}$ with compact support:

$$(1.14) \quad \begin{aligned} & \int_0^\infty du \int_0^\infty dh f(u, h) E(g(\tau_h^u)) \\ &= \int_0^\infty du E\left(\int_0^\infty d_t(L(t, u)) f(u, L(t, u)) g(t)\right) \\ &= \int_0^\infty du E\left(\int_0^\infty d_t(L(t, u)) g(t) E[f(u, L(t, u)) | B_t = u]\right) \\ &= \int_0^\infty du \int_0^\infty dt \frac{dE(L(t, u))}{dt} g(t) E[f(u, L(t, u)) | B_t = u] \\ &= \int_0^\infty du \int_0^\infty dt \frac{P(B_t \in du)}{du} g(t) E[f(u, L(t, u)) | B_t = u] \\ &= \int_0^\infty dt g(t) E[f(B_t, L(t, B_t))]. \end{aligned}$$

The first equality uses (1.2). The second equality follows from Fitzsimmons, Pitman and Yor (1993), Proposition 3. The fourth equality uses (1.13). \square

Next, abbreviate for $u, h \geq 0$,

$$(1.15) \quad \mathcal{P}_h^u = \left(\tau_h^u, \int_0^\infty L(\tau_h^u, u+v) dv, L(\tau_h^u, 0), \int_0^\infty L(\tau_h^u, -v) dv \right).$$

Then the distribution of \mathcal{P}_h^u is identified as in the following.

STEP 2. For every $u, h \geq 0$ and a.e. T, t_1, h_2, t_2 ,

$$(1.16) \quad \begin{aligned} & P(\mathcal{P}_h^u \in d(T, t_1, h_2, t_2)) \\ &= \varphi_h(t_1) \psi_{h, T-t_1-t_2}(u, h_2) \varphi_{h_2}(t_2) dT dt_1 dh_2 dt_2. \end{aligned}$$

PROOF. According to the RK theorems, $(L(\tau_h^u, -x))_{x \geq 0}$ is BESQ⁰ starting at $L(\tau_h^u, 0)$. Moreover, $L(\tau_h^u, 0)$ itself has distribution $\mathbb{P}_h(X_u)^{-1}$. Furthermore, from (0.2) we have

$$(1.17) \quad \tau_h^u = \int_0^\infty L(\tau_h^u, u+v) dv + \int_0^u L(\tau_h^u, u-v) dv + \int_0^\infty L(\tau_h^u, -v) dv.$$

Combining these statements with the RK theorems and (1.12), we obtain

$$(1.18) \quad \begin{aligned} P(\mathcal{J}_h^u \in d(T, t_1, h_2, t_2)) &= \mathbb{P}_h^* \left(\int_0^\infty X_v^* dv \in dt_1 \right) \mathbb{P}_{h_2}^* \left(\int_0^\infty X_v^* dv \in dt_2 \right) \\ &\quad \times \mathbb{P}_h \left(\int_0^u X_v dv \in d(T - t_1 - t_2), X_u \in dh_2 \right). \end{aligned}$$

But the r.h.s. of (1.18) equals the r.h.s. of (1.16), because of (1.10) and the identity $\{A(u) < T - t_1 - t_2\} = \{A^{-1}(T - t_1 - t_2) > u\}$ implied by (1.8). \square

STEP 3. $P(\tau_{L(T, B_T)}^{B_T} = T) = 1$.

PROOF. Simply note that $\tau_{L(T, B_T)}^{B_T} - T$ is distributed as the time change τ_0^0 for the process $(B_{T+t} - B_T)_{t \geq 0}$ [recall (1.2)]. But $P(\tau_0^0 = 0) = 1$ (see RY, Remark 1° following Proposition VI.2.5). \square

STEP 4. *Proof of Lemma 1.*

PROOF. First condition and integrate the l.h.s. of (1.11) w.r.t. the distribution of $(B_T, L(T, B_T))$, which is identified in Step 1. According to Step 3, we may then replace T by $\tau_{h_1}^u$ on $\{B_T = u, L(T, B_T) = h_1\}$. Next, condition and integrate w.r.t. the conditional distribution of $\mathcal{J}_{h_1}^u$ given $\{\tau_{h_1}^u = T\}$. Then the l.h.s. of (1.11) becomes

$$(1.19) \quad \begin{aligned} &\int_I du \int_0^\infty dh_1 \frac{P(\tau_{h_1}^u \in dT)}{dT} \int_{[0, \infty)^3} \frac{P(\mathcal{J}_{h_1}^u \in d(T, t_1, h_2, t_2))}{P(\tau_{h_1}^u \in dT)} \\ &\quad \times E \left(\Phi_1((L(\tau_{h_1}^u, u + x))_{x \geq 0}) \Phi_2((L(\tau_{h_1}^u, -x))_{x \geq 0}) \right. \\ &\quad \left. \times \Phi_3((L(\tau_{h_1}^u, u - x))_{x \in [0, u]}) \mid \mathcal{J}_{h_1}^u = (T, t_1, h_2, t_2) \right). \end{aligned}$$

Now use Step 2, apply the description of the local time processes provided by the RK theorems in combination with (1.12) and (1.15), and again use the elementary relation between A and A^{-1} stated at the end of the proof of Step 2. Then we obtain that (1.19) is equal to the r.h.s. of (1.11). \square

In Lemma 1, note that $A^*(\infty) = t_1$, respectively t_2 , corresponds to the Brownian motion spending t_1 , respectively t_2 , time units in the boundary areas $[B_T, \infty)$, respectively $(-\infty, 0]$, while $A^{-1}(T - t_1 - t_2)$ corresponds to the size of the middle area $[0, B_T]$ when the Brownian motion spends $T - t_1 - t_2$ time units there.

1.3. *Application to the Edwards model.* We are now ready to formulate the key representation of the expectation appearing in the l.h.s. of (1.1). This representation will be the starting point for the proof of Proposition 1 in Sections 2–4. Abbreviate

$$(1.20) \quad C_T = b^*T + C\sqrt{T}.$$

LEMMA 2. For all $T > 0$,

$$\begin{aligned}
 & E\left(\exp\left(-\int_{\mathbb{R}} L(T, x)^2 dx\right) 1_{0 < B_T \leq C_T}\right) \\
 &= \int_0^{C_T} du \int_{[0, \infty)^4} dt_1 dh_1 dt_2 dh_2 \\
 (1.21) \quad & \times \prod_{i=1}^2 \mathbb{E}_{h_i}^* \left(\exp\left(-\int_0^\infty X_v^{*2} dv\right) \middle| A^*(\infty) = t_i \right) \varphi_{h_i}(t_i) \\
 & \times \mathbb{E}_{h_1} \left(\exp\left(-\int_0^u X_v^2 dv\right) \middle| A^{-1}(T - t_1 - t_2) = u, X_u = h_2 \right) \\
 & \times \psi_{h_1, T-t_1-t_2}(u, h_2).
 \end{aligned}$$

The proof follows from Lemma 1.

Thus, we have expressed the expectation in the l.h.s. of (1.1) in terms of integrals over BESQ^0 and BESQ^2 and their additive functionals. Henceforth we can forget about the underlying Brownian motion and focus on these processes using their generators given in (1.3) and (1.4).

The importance of Lemma 2 is the decomposition into a *product* of three expectations. The main reason to introduce the densities φ_h and $\psi_{h_1, t}$ is the fact that the last factor in (1.21) depends on t_1 and t_2 . This dependence will vanish in the limit as $T \rightarrow \infty$, as we shall see in the sequel. After that the densities φ_h and $\psi_{h_1, t}$ can again be absorbed into the expectations [recall (1.10)].

2. A transformed Markov process. All we have done so far is to rewrite the key object of Proposition 1 in terms of expectations involving squared Bessel processes. We are now ready for our main attack.

In Section 2.1 we use Girsanov's formula to transform BESQ^2 into a new Markov process. The purpose of this transformation is to absorb the exponential factor appearing under the expectation in the fourth line of (1.21) into the transition probabilities of the new process. In Section 2.2 we list some properties of the transformed process. These are used in Section 2.3 to obtain a final reformulation of (1.21) on which the proof of Proposition 1 will be based. In Section 2.4 we formulate three main propositions, the proofs of which are deferred to Sections 3 and 4. In Section 2.5 the proof of Proposition 1 is completed subject to these propositions.

2.1. *Construction of the transformed process.* Fix $\alpha \in \mathbb{R}$ (later we shall pick $\alpha = \alpha^*$). Recall from Section 0.2 that $\rho(\alpha) \in \mathbb{R}$ is the largest eigenvalue of the operator \mathcal{H}^α defined in (0.5). We denote the corresponding strictly positive and L^2 -normalized eigenvector by x_α . From HH, Lemmas 20 and 22, we know that $x_\alpha: \mathbb{R}_0^+ \rightarrow \mathbb{R}^+$ is real-analytic with $\lim_{u \rightarrow \infty} u^{-3/2} \log x_\alpha(u) \in (-\infty, 0)$, and that $\alpha \mapsto x_\alpha \in L^2(\mathbb{R}_0^+)$ is real-analytic. Define

$$(2.1) \quad F_\alpha(u) = u^2 - \alpha u + \rho(\alpha), \quad u \in \mathbb{R}_0^+.$$

The following lemma defines the Girsanov transformation of BESQ² that we shall need later.

LEMMA 3. For $t, h_1, h_2 \geq 0$, let $P_t(h_1, dh_2)$ denote the transition probability function of BESQ². Then

$$(2.2) \quad \widehat{P}_t^a(h_1, dh_2) = \frac{x_a(h_2)}{x_a(h_1)} \mathbb{E}_{h_1} \left(\exp \left(- \int_0^t F_a(X_v) dv \right) \mid X_t = h_2 \right) P_t(h_1, dh_2)$$

defines the transition probability function of a diffusion $(X_v)_{v \geq 0}$ on \mathbb{R}_0^+ .

PROOF. Recall the definition of the generator G of BESQ² given in (1.4). According to RY, Section VIII.3, if $f \in C^2(\mathbb{R}_0^+)$ satisfies the equation

$$(2.3) \quad G(f) + \frac{1}{2}G(f^2) - fG(f) = F_a,$$

then

$$(2.4) \quad (D_t^{f, a})_{t \geq 0} = \left(\exp \left(f(X_t) - f(X_0) - \int_0^t F_a(X_s) ds \right) \right)_{t \geq 0}$$

is a local martingale under \mathbb{P}_h for any $h \geq 0$. Substitute $f = \log x$ in the l.h.s. of (2.3). Then an elementary calculation yields that for all $u \geq 0$,

$$(2.5) \quad \begin{aligned} \left(G(f) + \frac{1}{2}G(f^2) - fG(f) \right)(u) &= 2uf''(u) + 2f'(u) + 2uf'(u)^2 \\ &= \frac{2ux''(u) + 2x'(u)}{x(u)}. \end{aligned}$$

We now easily derive from the eigenvalue relation $\mathcal{H}_a x_a = \rho(a)x_a$ [recall (0.5)] that (2.3) is satisfied for $f = f_a = \log x_a$. Hence, $(D_t^{f_a, a})_{t \geq 0}$ is a local martingale under \mathbb{P}_h . Since F_a is bounded from below and x_a is bounded from above, each $D_t^{f_a, a}$ is bounded \mathbb{P}_h -a.s. Hence $(D_t^{f_a, a})_{t \geq 0}$ is a martingale under \mathbb{P}_h . The lemma now follows from RY, Proposition VIII.3.1. \square

We shall denote the distribution of the transformed process, conditioned on starting at $h \geq 0$, by $\widehat{\mathbb{P}}_h^a$ and the corresponding expectation by $\widehat{\mathbb{E}}_h^a$. Note that we have

$$(2.6) \quad \widehat{\mathbb{E}}_h^a(g(X_t)) = \mathbb{E}_h(D_t^{f_a, a} g(X_t)), \quad t \geq 0, g: \mathbb{R}_0^+ \rightarrow \mathbb{R}_0^+ \text{ measurable.}$$

2.2. Properties of the transformed process. We are going to list some properties of the process constructed in the preceding section.

1. The process introduced in Lemma 3 is a Feller process. According to RY, Proposition VIII.3.4, its generator is given by (recall $f_a = \log x_a$)

$$\begin{aligned}
 (2.7) \quad (\widehat{G}^a f)(u) &= (Gf)(u) + (G(f_a f) - f_a G(f) - f G(f_a))(u) \\
 &= (Gf)(u) + 4uf'_a(u)f'(u) \\
 &= 2uf''(u) + 2f'(u) \left(1 + 2u \frac{x'_a(u)}{x_a(u)} \right), \quad f \in C_c^2(\mathbb{R}^+).
 \end{aligned}$$

2. According to KS, Chapter 5, Equation (5.42), the scale function for the process is given (up to an affine transformation) by

$$(2.8) \quad s_a(u) = \int_c^u \frac{dv}{v x_a^2(v)}, \quad c > 0 \text{ arbitrary.}$$

Since x_a does not vanish at zero and has a subexponential tail at infinity (see the remarks at the beginning of Section 2.1), the scale function satisfies

$$(2.9) \quad \lim_{u \downarrow 0} s_a(u) = -\infty \quad \text{and} \quad \lim_{u \rightarrow \infty} s_a(u) = \infty.$$

3. The probability measure on \mathbb{R}_0^+ given by

$$(2.10) \quad \mu_a(du) = x_a(u)^2 du$$

is the normalized speed measure for the process [see KS, Chapter 5, Equation (5.51)]. Since it has finite mass, and because (2.9) holds, the process converges weakly towards μ_a from any starting point $h > 0$ (see KS, Chapter 5, Example 5.40), that is,

$$(2.11) \quad \lim_{t \rightarrow \infty} \widehat{\mathbb{E}}_h^a(f(X_t)) = \int_0^\infty f(u) \mu_a(du) \quad \text{for all bounded } f \in C(\mathbb{R}_0^+).$$

Using this convergence and the Feller property, one derives in a standard way that μ_a is the invariant distribution for the process. We write

$$(2.12) \quad \widehat{\mathbb{P}}^a = \int_0^\infty \widehat{\mathbb{P}}_h^a \mu_a(dh)$$

to denote the distribution of the process starting in the invariant distribution and write $\widehat{\mathbb{E}}^a$ for the corresponding expectation.

4. According to Ethier and Kurtz (1986), Theorem 6.1.4, the process $(Y_t)_{t \geq 0}$ given by

$$(2.13) \quad Y_t = X_{A^{-1}(t)}, \quad t \geq 0$$

is a diffusion under $\widehat{\mathbb{P}}^a$ with generator

$$(2.14) \quad (\widetilde{G}^a f)(u) = \frac{1}{u} (\widehat{G}^a f)(u), \quad u > 0, \quad f \in C_c^2(\mathbb{R}^+)$$

[see (2.7)]. This process has the same scale function s_a as $(X_t)_{t \geq 0}$ [see (2.8)], and its normalized speed measure is given by

$$(2.15) \quad \nu_a(du) = \frac{u}{\rho'(a)} x_a^2(u) du.$$

[In order to see that $\nu_a(\mathbb{R}^+) = 1$, differentiate the relation $\rho(a) = \langle x_a, \mathcal{H}^a x_a \rangle_{L^2}$ w.r.t. a . Use (0.5) and the relation $(d/da)\langle x_a, x_a \rangle_{L^2} = 0$.] Similarly as in (2.11), for any starting point $h > 0$,

$$(2.16) \quad \lim_{t \rightarrow \infty} \widehat{\mathbb{E}}_h^a(f(Y_t)) = \int_0^\infty f(u) \nu_a(du) \quad \text{for all bounded } f \in C(\mathbb{R}_0^+)$$

and hence ν_a is the invariant distribution of the process $(Y_t)_{t \geq 0}$. We write

$$(2.17) \quad \widetilde{\mathbb{P}}^a = \int_0^\infty \widehat{\mathbb{P}}_h^a \nu_a(dh)$$

to denote the distribution of the process $(X_t)_{t \geq 0}$ starting in the invariant distribution ν_a of the process $(Y_t)_{t \geq 0}$ and we write $\widehat{\mathbb{E}}^a$ for the corresponding expectation.

2.3. *Final reformulation.* Using the representation in Lemma 2, we shall rewrite the l.h.s. of (1.1) in terms of the transformed process introduced in Lemma 3. This will be the final reformulation in terms of which the proof of Proposition 1 will be finished in Sections 2.4–2.5.

For $h, t \geq 0$ and $a \in \mathbb{R}$, introduce the abbreviation [recall (1.9) and (1.10)]

$$(2.18) \quad \begin{aligned} F_a^*(u) &= -u^2 + au, \quad u \in \mathbb{R}_0^+, \\ w_a(h, t) &= \mathbb{E}_h^* \left(\exp \left(- \int_0^\infty F_a^*(X_v^*) dv \right) \mid A^*(\infty) = t \right) \varphi_h(t) \\ &= \exp(at) w_0(h, t). \end{aligned}$$

Recall that $\widehat{\mathbb{E}}^a$ denotes the expectation for the transformed process $(X_t)_{t \geq 0}$, starting in the invariant starting distribution μ_a given by (2.10).

LEMMA 4. For every $T > 0$,

$$(2.19) \quad \begin{aligned} & \exp(a^*T) E \left(\exp \left(- \int_{\mathbb{R}} L(T, x)^2 dx \right) 1_{0 < B_T \leq C_T} \right) \\ &= \int_0^\infty dt_1 \int_0^\infty dt_2 \\ & \quad \times \widehat{\mathbb{E}}^{a^*} \left(\frac{w_{a^*}(X_0, t_1)}{x_{a^*}(X_0)} 1_{A^{-1}(T-t_1-t_2) \leq C_T} \frac{w_{a^*}(X_{A^{-1}(T-t_1-t_2)}, t_2)}{x_{a^*}(X_{A^{-1}(T-t_1-t_2)})} \right). \end{aligned}$$

PROOF. First, from (1.8), (2.1) and $\rho(a^*) = 0$ it follows that on $\{A^{-1}(t) = u\}$,

$$(2.20) \quad a^*t - \int_0^u X_v^2 dv = - \int_0^u F_{a^*}(X_v) dv, \quad t, u \geq 0.$$

By an absolute continuous transformation from \mathbb{P}_h to $\widehat{\mathbb{P}}_h^{\alpha^*}$, we therefore obtain via (2.2) the identity [recall (1.10)]

$$(2.21) \quad \begin{aligned} & \exp(\alpha^* t) \mathbb{E}_{h_1} \left(\exp \left(- \int_0^u X_v^2 dv \right) \middle| A^{-1}(t) = u, X_u = h_2 \right) \psi_{h_1, t}(u, h_2) du dh_2 \\ &= \widehat{\mathbb{P}}_{h_1}^{\alpha^*} (A^{-1}(t) \in du, X_u \in dh_2) \frac{x_{\alpha^*}(h_1)}{x_{\alpha^*}(h_2)} \end{aligned}$$

for a.e. $u, h_1, h_2, t \geq 0$. Similarly to (2.20), we have on $\{\int_0^\infty X_v^* dv = t\}$,

$$(2.22) \quad \alpha^* t - \int_0^\infty (X_v^*)^2 dv = - \int_0^\infty F_{\alpha^*}(X_v^*) dv, \quad t \geq 0$$

and hence

$$(2.23) \quad \begin{aligned} & \exp(\alpha^* t_i) \mathbb{E}_{h_i}^* \left(\exp \left(- \int_0^\infty (X_v^*)^2 dv \right) \middle| A^*(\infty) = t_i \right) \varphi_{h_i}(t_i) \\ &= w_{\alpha^*}(h_i, t_i), \quad i = 1, 2. \end{aligned}$$

Next, note that the l.h.s. of (2.19) is equal to the l.h.s. of (1.21) times the factor $e^{\alpha^* T}$. We divide this factor into three parts, according to the identity $T = t_1 + (T - t_1 - t_2) + t_2$, and assign them to each of the three expectations in the r.h.s. of (1.21). Substitute (2.21) with $t = T - t_1 - t_2$ and (2.23) into (1.21). Then we obtain that the l.h.s. of (2.19) is equal to

$$(2.24) \quad \begin{aligned} & \int_{[0, \infty)^4} dh_1 dh_2 dt_1 dt_2 w_{\alpha^*}(h_1, t_1) w_{\alpha^*}(h_2, t_2) \frac{x_{\alpha^*}(h_1)}{x_{\alpha^*}(h_2)} \\ & \times \widehat{\mathbb{P}}_{h_1}^{\alpha^*} (A^{-1}(T - t_1 - t_2) \leq C_T, X_{A^{-1}(T - t_1 - t_2)} \in dh_2). \end{aligned}$$

Now formally carry out the integration over h_1, h_2 , recalling (2.10) and (2.12), to arrive at the r.h.s. of (2.19). \square

Roughly speaking, the function w_{α^*} in the r.h.s. of (2.19) describes the contribution to the random variable $\exp[-\int_{\mathbb{R}} L(T, x)^2 dx]$ coming from the boundary pieces [i.e., the parts of the path in $(-\infty, 0] \cup [B_T, \infty)$], while A^{-1} gives the size of the area over which the middle piece (i.e., the parts of the path in $[0, B_T]$) spreads out.

2.4. Key steps in the proof of Proposition 1. The proof of Proposition 1 now basically requires the following three ingredients.

1. A CLT for $(A^{-1}(t))_{t \geq 0}$ under $\widehat{\mathbb{P}}^{\alpha^*}$.
2. An extension of the weak convergence of $(Y_t)_{t \geq 0} = (X_{A^{-1}(t)})_{t \geq 0}$ stated in (2.16).
3. Some integrability properties of w_{α^*} .

The precise statements that we shall need are formulated in Propositions 2–4. The proof of these propositions is deferred to Sections 3 and 4.

We need some more notation. Let $\langle \cdot, \cdot \rangle_{L^2}$ denote the standard inner product on $L^2(\mathbb{R}_0^+)$. Let $\langle \cdot, \cdot \rangle_{L^2}^\circ$ denote the weighted inner product

$$(2.25) \quad \langle f, g \rangle_{L^2}^\circ = \int_0^\infty dh hf(h)g(h)$$

on $L^{2,\circ}(\mathbb{R}_0^+) = \{f: \mathbb{R}_0^+ \rightarrow \mathbb{R} \text{ measurable} \mid \int_0^\infty dh hf^2(h) < \infty\}$. We write $\|\cdot\|_{L^2}$, respectively, $\|\cdot\|_{L^2}^\circ$ for the corresponding norms.

For bounded and measurable $f, g: \mathbb{R}_0^+ \rightarrow \mathbb{R}$, $T \geq 0$ and $a \in \mathbb{R}$, abbreviate [recall Lemma 3, (2.10) and (2.12)]

$$(2.26) \quad \begin{aligned} N_{T,a}^{f,g} &= \widehat{\mathbb{E}}^a \left(\frac{f}{x_a}(Y_0) \frac{g}{x_a}(Y_T) \right) \\ &= \int_0^\infty dh f(h) \mathbb{E}_h \left(\exp \left(- \int_0^{A^{-1}(T)} F_a(X_s) ds \right) g(X_{A^{-1}(T)}) \right). \end{aligned}$$

Furthermore, define

$$(2.27) \quad \sigma^2(a) = \frac{\rho''(a)}{\rho'(a)^3}$$

and note that $\sigma^2(a^*) = c^{*2}$ as defined in (0.6). Denote by $\rho^{-1}: \mathbb{R} \rightarrow \mathbb{R}$ the inverse function of $\rho: \mathbb{R} \rightarrow \mathbb{R}$.

PROPOSITION 2. *For all bounded and measurable $f, g: \mathbb{R}_0^+ \rightarrow \mathbb{R}$, all $a, \lambda \in \mathbb{R}$ and all $T, T' \geq 0$,*

$$(2.28) \quad \begin{aligned} &\widehat{\mathbb{E}}^a \left(\frac{f}{x_a}(Y_0) \exp \left(\frac{\lambda}{\sqrt{T}} \left(A^{-1}(T') - \frac{T}{\rho'(a)} \right) \right) \frac{g}{x_a}(Y_{T'}) \right) \\ &= \exp \left(\frac{\lambda^2}{2} \sigma^2(\xi_T) \right) N_{T',a_{\lambda,T}}^{f,g} \exp((T - T')(a_{\lambda,T} - a)), \end{aligned}$$

where

$$(2.29) \quad a_{\lambda,T} = \rho^{-1} \left(\rho(a) - \frac{\lambda}{\sqrt{T}} \right)$$

and $\xi_T \in [a, a_{\lambda,T}] \cup [a_{\lambda,T}, a]$.

PROPOSITION 3. *Let $f, g: \mathbb{R}_0^+ \rightarrow \mathbb{R}$ be measurable such that $f/\text{id}, g \in L^{2,\circ}$. Then for every $a \in \mathbb{R}$ and $a_T \rightarrow a$,*

$$(2.30) \quad \lim_{T \rightarrow \infty} N_{T,a_T}^{f,g} = \frac{1}{\rho'(a)} \langle f, x_a \rangle_{L^2} \langle g, x_a \rangle_{L^2}^\circ.$$

Next, recall (2.18). For $a \in \mathbb{R}$, define $y_a: \mathbb{R}_0^+ \rightarrow [0, \infty]$ by

$$(2.31) \quad y_a(h) = \int_0^\infty w_a(h, t) dt = \mathbb{E}_h^* \left(\exp \left(- \int_0^\infty F_a^*(X_v^*) dv \right) \right).$$

Furthermore, for $p \in (1, 2)$, $q \in (2, \infty)$ and $t > 0$ define

$$(2.32) \quad \begin{aligned} W_p^{(1)}(t) &= \left(\int_0^\infty h^{1-p} x_{a^*}(h)^{2-p} w_{a^*}(h, t)^p dh \right)^{1/p}, \\ W_q^{(2)}(t) &= \left(\int_0^\infty h x_{a^*}(h)^{2-q} w_{a^*}(h, t)^q dh \right)^{1/q}. \end{aligned}$$

PROPOSITION 4.

- (i) y_{a^*} is measurable and bounded.
- (ii) $\varphi_h(t) = (h/2\sqrt{2\pi t^3}) \exp(-h^2/8t)$ for all $h \geq 0$ and $t > 0$.
- (iii) For any $p \in (1, 2)$, $W_p^{(1)}$ is integrable on \mathbb{R}^+ .
- (iv) For any $q \in (2, \infty)$ sufficiently close to 2, $W_q^{(2)}$ is integrable on \mathbb{R}^+ .

2.5. *Proof of Proposition 1.* In this subsection we complete the proof of Proposition 1, subject to Propositions 2–4. We shall show that (1.1) follows from (2.19), with S identified as

$$(2.33) \quad S = b^* \langle y_{a^*}, x_{a^*} \rangle_{L^2} \langle y_{a^*}, x_{a^*} \rangle_{L^2}^\circ.$$

STEP 1. For all $t_1, t_2 > 0$, as $T \rightarrow \infty$, the integrand on the r.h.s. of (2.19) tends to

$$b^* \langle w_{a^*}(\cdot, t_1), x_{a^*} \rangle_{L^2} \langle w_{a^*}(\cdot, t_2), x_{a^*} \rangle_{L^2}^\circ \mathcal{N}_{c^{*2}}((-\infty, C]).$$

PROOF. By Proposition 4(ii) and (2.18), the functions $f = w_{a^*}(\cdot, t_1)$ and $g = w_{a^*}(\cdot, t_2)$ satisfy the assumptions of Proposition 3 for all $t_1, t_2 > 0$, since they are bounded by a factor times $\varphi(\cdot, t_1)$, respectively, $\varphi(\cdot, t_2)$. Define a (non-Markovian) path measure $\mathbb{P}_{T,a}^{f,g}$ by

$$(2.34) \quad \frac{d\mathbb{P}_{T,a}^{f,g}}{d\widehat{\mathbb{P}}^a} = \frac{1}{N_{T,a}^{f,g}} \frac{f}{x_a}(Y_0) \frac{g}{x_a}(Y_T).$$

Write $\mathbb{E}_{T,a}^{f,g}$ for the corresponding expectation. Apply Proposition 2 for $a = a^*$ and $T' = T - t_1 - t_2$ to obtain that for every $\lambda \in \mathbb{R}$ and $T \geq t_1 + t_2$,

$$(2.35) \quad \begin{aligned} &\mathbb{E}_{T-t_1-t_2, a^*}^{f,g} \left(\exp \left(\frac{\lambda}{\sqrt{T}} [A^{-1}(T - t_1 - t_2) - b^*T] \right) \right) \\ &= \exp \left(\frac{\lambda^2}{2} \sigma^2(\xi_T^*) \right) \frac{N_{T-t_1-t_2, a_{\lambda, T}^*}^{f,g}}{N_{T-t_1-t_2, a^*}^{f,g}} \exp((t_1 + t_2)(a_{\lambda, T}^* - a^*)), \end{aligned}$$

where $\rho(a^*) = 0$, $b^* = 1/\rho'(a^*)$ [recall (0.6)], $a_{\lambda, T}^* = \rho^{-1}(-\lambda/\sqrt{T})$ and $\xi_T^* \in [a^*, a_{\lambda, T}^*] \cup [a_{\lambda, T}^*, a^*]$. Since ρ' , ρ'' and ρ^{-1} are continuous, we have $a_{\lambda, T}^* \rightarrow a^*$ and $\sigma^2(\xi_T^*) \rightarrow c^{*2}$ as $T \rightarrow \infty$. Therefore, by Proposition 3, the r.h.s. of (2.35) tends to $\exp((\lambda^2/2)c^{*2})$ as $T \rightarrow \infty$. Thus, the distribution of $(1/\sqrt{T})[A^{-1}(T -$

$t_1 - t_2) - b^*T]$ under $\mathbb{P}_{T-t_1-t_2, a^*}^{f, g}$ converges weakly towards \mathcal{N}_{c^*2} . Via (2.34), this in turn implies that [recall (1.20)]

$$\begin{aligned}
 (2.36) \quad & \lim_{T \rightarrow \infty} \widehat{\mathbb{E}}^{a^*} \left(\frac{w_{a^*}(X_0, t_1)}{x_{a^*}(X_0)} 1_{A^{-1}(T-t_1-t_2) \leq C_T} \frac{w_{a^*}(X_{A^{-1}(T-t_1-t_2)}, t_2)}{x_{a^*}(X_{A^{-1}(T-t_1-t_2)})} \right) \\
 &= \lim_{T \rightarrow \infty} N_{T-t_1-t_2, a^*}^{f, g} \mathbb{P}_{T-t_1-t_2, a^*}^{f, g} (A^{-1}(T-t_1-t_2) - b^*T \leq C\sqrt{T}) \\
 &= b^* \langle f, x_{a^*} \rangle_{L^2} \langle g, x_{a^*} \rangle_{L^2} \mathcal{N}_{c^*2}((-\infty, C]),
 \end{aligned}$$

again according to Proposition 3. \square

STEP 2. For all $t_1, t_2 > 0$, and any $p, q > 1$ satisfying $1/p + 1/q = 1$, the integrand on the r.h.s. of (2.19) is bounded uniformly in $T > 0$ by $W_p^{(1)}(t_1)W_q^{(2)}(t_2)$ defined in (2.32).

PROOF. Recall (3) and (4) in Section 2.2. Make a change of measure from $\widehat{\mathbb{E}}^{a^*}$ to $\widetilde{\mathbb{E}}^{a^*}$, use the Hölder inequality and the stationarity of $(Y_t)_{t \geq 0}$ under $\widetilde{\mathbb{P}}^{a^*}$ [recall (2.15) and (2.17)], to obtain

$$\begin{aligned}
 (2.37) \quad & \widehat{\mathbb{E}}^{a^*} \left(\frac{w_{a^*}(X_0, t_1)}{x_{a^*}(X_0)} 1_{A^{-1}(T-t_1-t_2) \leq C_T} \frac{w_{a^*}(X_{A^{-1}(T-t_1-t_2)}, t_2)}{x_{a^*}(X_{A^{-1}(T-t_1-t_2)})} \right) \\
 & \leq \rho'(a^*) \widetilde{\mathbb{E}}^{a^*} \left(\frac{w_{a^*}(Y_0, t_1)}{Y_0 x_{a^*}(Y_0)} \frac{w_{a^*}(Y_{T-t_1-t_2}, t_2)}{x_{a^*}(Y_{T-t_1-t_2})} \right) \\
 & \leq \rho'(a^*) \left(\widetilde{\mathbb{E}}^{a^*} \left(\left[\frac{w_{a^*}(Y_0, t_1)}{Y_0 x_{a^*}(Y_0)} \right]^p \right) \right)^{1/p} \left(\widetilde{\mathbb{E}}^{a^*} \left(\left[\frac{w_{a^*}(Y_{T-t_1-t_2}, t_2)}{x_{a^*}(Y_{T-t_1-t_2})} \right]^q \right) \right)^{1/q} \\
 & = W_p^{(1)}(t_1) W_q^{(2)}(t_2). \quad \square
 \end{aligned}$$

STEP 3. Conclusion of the proof.

PROOF. Let $T \rightarrow \infty$ in (2.19) and note that, for some $p, q > 1$ satisfying $1/p + 1/q = 1$, the bound in Step 2 is integrable in $(t_1, t_2) \in (\mathbb{R}^+)^2$ by Proposition 4(iii) and 4(iv). Therefore, by Steps 1 and 2 and the dominated convergence theorem we may interchange $T \rightarrow \infty$ and $\int_0^\infty dt_1 \int_0^\infty dt_2$, to obtain

$$\begin{aligned}
 (2.38) \quad & \lim_{T \rightarrow \infty} \text{l.h.s. of (2.19)} \\
 &= b^* \int_0^\infty dt_1 \int_0^\infty dt_2 \langle w_{a^*}(\cdot, t_1), x_{a^*} \rangle_{L^2} \langle w_{a^*}(\cdot, t_2), x_{a^*} \rangle_{L^2} \mathcal{N}_{c^*2}((-\infty, C]).
 \end{aligned}$$

Now use (2.31), Fubini's theorem and Proposition 4(i) to identify the r.h.s. of (2.38) as $S \mathcal{N}_{c^*2}((-\infty, C])$, with S given in (2.33). \square

3. CLT for the middle piece. This section contains the proofs of Propositions 2 and 3.

3.1. *Proof of Proposition 2.* Recall Lemma 3 and (2.26) to see that the l.h.s. of (2.28) is equal to

$$(3.1) \quad \exp\left(-\frac{\lambda\sqrt{T}}{\rho'(a)}\right) \int_0^\infty dh f(h) \\ \times \mathbb{E}_h\left(\exp\left(-\int_0^{A^{-1}(T')} \left(F_a(X_s) - \frac{\lambda}{\sqrt{T}}\right) ds\right) g(X_{A^{-1}(T')})\right).$$

According to (2.29), $\rho(a_{\lambda,T}) = \rho(a) - (\lambda/\sqrt{T})$. Since $T' = \int_0^{A^{-1}(T')} X_s ds$ [see (1.8)] and $F_a(u) = u^2 - au + \rho(a)$ [see (2.1)], we may write the exponents in (3.1) as

$$(3.2) \quad -\int_0^{A^{-1}(T')} F_{a_{\lambda,T}}(X_s) ds + (a - a_{\lambda,T}) \int_0^{A^{-1}(T')} X_s ds - \frac{\lambda\sqrt{T}}{\rho'(a)} \\ = -\int_0^{A^{-1}(T')} F_{a_{\lambda,T}}(X_s) ds + T\left(a - a_{\lambda,T} - \frac{\lambda}{\sqrt{T}\rho'(a)}\right) \\ + (T - T')(a_{\lambda,T} - a).$$

Substitute this into (3.1) and use (2.26) to get that

$$(3.3) \quad \text{l.h.s. of (2.28)} \\ = \exp\left(T\left(a - a_{\lambda,T} - \frac{\lambda}{\sqrt{T}\rho'(a)}\right)\right) N_{T', a_{\lambda,T}}^{f,g} \exp((T - T')(a_{\lambda,T} - a)).$$

Next, expand the inverse function ρ^{-1} of ρ as a Taylor series around $\rho(a)$ up to second order. It follows that there is an r_T between $\rho(a)$ and $\rho(a) - (\lambda/\sqrt{T})$ such that

$$(3.4) \quad a_{\lambda,T} = \rho^{-1}\left(\rho(a) - \frac{\lambda}{\sqrt{T}}\right) = \rho^{-1}(\rho(a)) - \frac{\lambda}{\sqrt{T}}(\rho^{-1})'(\rho(a)) + \frac{\lambda^2}{2T}(\rho^{-1})''(r_T) \\ = a - \frac{\lambda}{\sqrt{T}\rho'(a)} - \frac{\lambda^2}{2T} \frac{\rho''}{(\rho')^3}(\rho^{-1}(r_T)) = a - \frac{\lambda}{\sqrt{T}\rho'(a)} - \frac{\lambda^2}{2T} \sigma^2(\xi_T)$$

[see (2.27)] with $\xi_T = \rho^{-1}(r_T)$. Observe that ξ_T is between a and $a_{\lambda,T}$ by the monotonicity of ρ . Now substitute (3.4) into (3.3) to arrive at (2.28). \square

3.2. *Proof of Proposition 3.* We shall use an expansion in terms of the eigenfunctions of the operator $\mathcal{M}^a: L^2, \circ(\mathbb{R}_0^+) \cap C^2(\mathbb{R}_0^+) \rightarrow C(\mathbb{R}_0^+)$ defined by

$$(3.5) \quad (\mathcal{M}^a x)(u) = \frac{(\mathcal{K}^a x)(u) - \rho(a)x(u)}{u}$$

[recall (0.5)]. Obviously, \mathcal{M}^a is symmetric w.r.t. $\langle \cdot, \cdot \rangle_{L^2}^\circ$ because \mathcal{K}^a is symmetric w.r.t. $\langle \cdot, \cdot \rangle_{L^2}$. Also \mathcal{M}^a is a Sturm–Liouville operator. We are going to identify its eigenvalues and eigenvectors in terms of those of \mathcal{K}^a .

For $l \in \mathbb{N}_0$, let $\rho^{(l)}(a)$ denote the l th largest eigenvalue of \mathcal{K}^a and $x_a^{(l)} \in L^2(\mathbb{R}^+)$ the corresponding eigenfunction, normalized such that $\|x_a^{(l)}\|_{L^2} = 1$

(all eigenspaces are one-dimensional by HH, Lemma 20). Then $\rho^{(0)} = \rho$, and each $\rho^{(l)}$ is continuous and strictly increasing [differentiate the formula $\rho^{(l)}(a) = \langle x_a^{(l)}, \mathcal{H}^a x_a^{(l)} \rangle_{L^2}$ to obtain $d/da\rho^{(l)}(a) = \|x_a^{(l)}\|_{L^2}^{\circ 2}$ via (0.5)]. Moreover, $\lim_{a \rightarrow \pm\infty} \rho^{(l)}(a) = \pm\infty$. Since $x_a^{(l)}$ has a subexponentially small tail at infinity (see HH, Lemma 20), it is also an element of $L^{2,\circ}(\mathbb{R}_0^+)$.

Next, define $\alpha^{(l)}(a) \in \mathbb{R}$ and $y_a^{(l)} \in L^{2,\circ}(\mathbb{R}_0^+)$ by

$$(3.6) \quad \rho^{(l)}(a - \alpha^{(l)}(a)) = \rho(a) \quad \text{and} \quad y_a^{(l)} = \frac{x_{a-\alpha^{(l)}(a)}^{(l)}}{\|x_{a-\alpha^{(l)}(a)}^{(l)}\|_{L^2}^{\circ}}, \quad l \in \mathbb{N}_0.$$

Note that $\alpha^{(0)}(a) = 0$, $y_a^{(0)} = x_a/\sqrt{\rho'(a)}$, and $\alpha^{(l+1)}(a) < \alpha^{(l)}(a)$ for all $l \in \mathbb{N}_0$ since $\rho^{(l)}(a)$ is strictly decreasing in l and strictly increasing in a .

STEP 1. For each $a \in \mathbb{R}$, the sequence $(y_a^{(l)})_{l \in \mathbb{N}_0}$ is an orthonormal basis in $L^{2,\circ}(\mathbb{R}^+)$.

PROOF. Since \mathcal{M}^a is a symmetric Sturm–Liouville operator, all its eigenspaces are orthogonal to each other and one-dimensional, and they span the space $L^{2,\circ}(\mathbb{R}^+)$. Thus, it suffices to show that the functions $y_a^{(0)}, y_a^{(1)}, \dots$ are all the eigenfunctions of \mathcal{M}^a . Now, from (0.5) and (3.5) we easily derive the equivalence

$$(3.7) \quad \mathcal{M}^a x = \alpha x \iff \mathcal{H}^{a-\alpha} x = \rho(a)x,$$

which is valid for every $a, \alpha \in \mathbb{R}$ and $x \in C^2(\mathbb{R}_0^+)$. From (3.6) and (3.7) we see that $(\alpha^{(l)}(a))_{l \in \mathbb{N}_0}$ is the sequence of all the eigenvalues of \mathcal{M}^a with corresponding eigenfunctions $(y_a^{(l)})_{l \in \mathbb{N}_0}$, since (3.7) implies that for every eigenvalue α of \mathcal{M}^a there is an $l \in \mathbb{N}_0$ such that $\rho^{(l)}(a - \alpha) = \rho(a)$. \square

STEP 2. For every $h, T \geq 0$, $l \in \mathbb{N}_0$ and $a \in \mathbb{R}$,

$$(3.8) \quad \widehat{\mathbb{E}}_h^a \left(\frac{y_a^{(l)}}{x_a} (Y_T) \right) = \exp(\alpha^{(l)}(a)T) \frac{y_a^{(l)}}{x_a}(h).$$

PROOF. Use (2.7) and (2.14) to compute, for $f \in C^2(\mathbb{R}^+)$,

$$(3.9) \quad \left(\tilde{G}^a \left(\frac{f}{x_a} \right) \right) (u) = \frac{f(u)}{ux_a(u)} \left(\frac{2uf''(u) + 2f'(u)}{f(u)} - \frac{2ux_a''(u) + 2x_a'(u)}{x_a(u)} \right).$$

Apply this for $f = y_a^{(l)}$, use (0.5) and the eigenvalue relation $\mathcal{H}^{a'} x_{a'}^{(l)} = \rho^{(l)}(a') x_{a'}^{(l)}$ for $(a', l) = (a, 0)$ and for $(a', l) = (a - \alpha^{(l)}(a), l)$, to obtain

$$(3.10) \quad \tilde{G}^a \left(\frac{y_a^{(l)}}{x_a} \right) = \alpha^{(l)}(a) \frac{y_a^{(l)}}{x_a}.$$

Thus, \tilde{G}^a being the generator of the process $(Y_t)_{t \geq 0}$, the function $f(T) = \widehat{\mathbb{E}}_h^a((y_a^{(l)}/x_a)(Y_T))$ satisfies the differential equation $f' = \alpha^{(l)}(a)f$. Therefore $f(T) = \exp(\alpha^{(l)}(a)T)f(0)$, which is our assertion. \square

STEP 3. *Conclusion of the proof.*

PROOF. According to Step 1, we may expand $g \in L^{2,\circ}(\mathbb{R}_0^+)$ as

$$(3.11) \quad \begin{aligned} g &= \sum_{l=0}^{\infty} y_{a_T}^{(l)} \langle g, y_{a_T}^{(l)} \rangle_{L^2}^{\circ} \\ &= \frac{x_{a_T}}{\rho'(a_T)} \langle g, x_{a_T} \rangle_{L^2}^{\circ} + \sum_{l=1}^{\infty} y_{a_T}^{(l)} \langle g, y_{a_T}^{(l)} \rangle_{L^2}^{\circ} \quad (T \geq 0). \end{aligned}$$

Substitute this into (2.26) to obtain [recall (2.10) and (2.12)]

$$(3.12) \quad \begin{aligned} & \left| N_{T,a_T}^{f,g} - \frac{1}{\rho'(a)} \langle f, x_a \rangle_{L^2} \langle g, x_a \rangle_{L^2}^{\circ} \right| \\ & \leq \left| \frac{1}{\rho'(a_T)} \langle f, x_{a_T} \rangle_{L^2} \langle g, x_{a_T} \rangle_{L^2}^{\circ} - \frac{1}{\rho'(a)} \langle f, x_a \rangle_{L^2} \langle g, x_a \rangle_{L^2}^{\circ} \right| \\ & \quad + \sum_{l=1}^{\infty} \left| \left(\int_0^{\infty} dh f(h) x_{a_T}(h) \widehat{\mathbb{E}}_h^a \left(\frac{y_{a_T}^{(l)}}{x_{a_T}}(Y_T) \right) \right) \langle g, y_{a_T}^{(l)} \rangle_{L^2}^{\circ} \right|. \end{aligned}$$

With the help of Step 2, the second term on the r.h.s. of (3.12) equals

$$(3.13) \quad \begin{aligned} & \sum_{l=1}^{\infty} \exp(\alpha^{(l)}(a_T)T) \left| \left(\int_0^{\infty} dh f(h) x_{a_T}(h) \frac{y_{a_T}^{(l)}}{x_{a_T}}(h) \right) \langle g, y_{a_T}^{(l)} \rangle_{L^2}^{\circ} \right| \\ & \leq \exp(\alpha^{(1)}(a_T)T) \sum_{l=0}^{\infty} \left| \left\langle \frac{f}{\text{id}}, y_{a_T}^{(l)} \right\rangle_{L^2}^{\circ} \langle g, y_{a_T}^{(l)} \rangle_{L^2}^{\circ} \right| \\ & \leq \exp(\alpha^{(1)}(a_T)T) \sqrt{\sum_{l=0}^{\infty} \left(\left\langle \frac{f}{\text{id}}, y_{a_T}^{(l)} \right\rangle_{L^2}^{\circ} \right)^2} \sqrt{\sum_{l=0}^{\infty} \left(\langle g, y_{a_T}^{(l)} \rangle_{L^2}^{\circ} \right)^2} \\ & = \exp(\alpha^{(1)}(a_T)T) \left\| \frac{f}{\text{id}} \right\|_{L^2}^{\circ} \|g\|_{L^2}^{\circ}. \end{aligned}$$

This tends to zero as $T \rightarrow \infty$ since $\lim_{T \rightarrow \infty} \alpha^{(1)}(a_T) = \alpha^{(1)}(a) < 0$. The first term on the r.h.s. of (3.12) vanishes as $T \rightarrow \infty$ because of the continuity of $a \mapsto x_a \in L^2(\mathbb{R}^+)$ and $a \mapsto \rho'(a)$ (see HH, Lemma 22). \square

4. Integrability for the boundary pieces. This section contains the proof of Proposition 4. It turns out that the functions w_a [in (2.18)] and y_a [in (2.31)] have a nice representation in terms of standard one-dimensional Brownian motion and that y_a is a transformation of the Airy function. This will be explored in Section 4.2. Section 4.1 contains some preparations.

4.1. *Preparations.* Let $\text{Ai}: \mathbb{R} \rightarrow \mathbb{R}$ denote the Airy function, that is, the unique (modulo a constant multiple) solution of the Airy equation

$$(4.1) \quad x''(u) - ux(u) = 0, \quad u \in \mathbb{R}$$

that is bounded on \mathbb{R}_0^+ . Let $u_1 = \sup\{u \in \mathbb{R} \mid \text{Ai}(u) = 0\}$ be its largest zero. From Abramowitz and Stegun (1970), Table 10.13 and page 450, it is known that $u_1 = -2.3381\dots$. For $a < -2^{1/3}u_1$, define $z_a: \mathbb{R}_0^+ \rightarrow \mathbb{R}^+$ by

$$(4.2) \quad z_a(u) = \frac{\text{Ai}(2^{-1/3}(u - a))}{\text{Ai}(-2^{-1/3}a)}, \quad u \geq 0.$$

In Lemma 8 in Section 4.2, z_a will turn out to be equal to y_a . Some of its properties are given in the following lemma.

LEMMA 5. *For all $a < -2^{1/3}u_1$, the function z_a is real-analytic, strictly positive on \mathbb{R}_0^+ with $z_a(0) = 1$, and satisfies*

$$(4.3) \quad 2z_a''(u) + (a - u)z_a(u) = 0, \quad u \geq 0.$$

Moreover,

$$(4.4) \quad \lim_{u \rightarrow \infty} u^{-3/2} \log z_a(u) \in (-\infty, 0).$$

PROOF. It is well known that Ai is analytic. From (4.2) and the definition of u_1 , it is clear that $z_a(0) = 1$ and that $z_a(u) > 0$ for $u \geq 0$. Equation (4.3) follows easily from (4.1). The asymptotics in (4.4) follows from Abramowitz and Stegun (1970), 10.4.59. \square

The following lemma shows in particular that Lemma 5 can be used for $a = a^*$.

$$\text{LEMMA 6. } a^* \leq \frac{3}{2}\pi^{1/3} < -u_1.$$

PROOF. The first inequality is proved via the variational representation

$$(4.5) \quad a^* = \inf_{x \in L^2(\mathbb{R}_0^+) \cap C^2(\mathbb{R}_0^+): x \neq 0} \frac{\int_0^\infty [u^2 x^2(u) + 2ux'(u)^2] du}{\int_0^\infty ux^2(u) du}.$$

This representation stems from the relation (see HH, Section 5.1)

$$(4.6) \quad 0 = \rho(a^*) = \max_{x \in L^2(\mathbb{R}_0^+) \cap C^2(\mathbb{R}_0^+): \|x\|_{L^2} = 1} \langle x, \mathcal{H}^{a^*} x \rangle_{L^2},$$

in which, by (0.5),

$$(4.7) \quad \langle x, \mathcal{H}^{a^*} x \rangle_{L^2} = \int_0^\infty [(a^*u - u^2)x(u)^2 - 2ux'(u)^2] du.$$

In (4.5), we choose the test function

$$(4.8) \quad x(u) = \exp\left(-u^2 \frac{\pi^{1/3}}{8}\right).$$

Elementary computations give that $\int_0^\infty ux^2(u) du = 2\pi^{-1/3}$ and $\int_0^\infty u^2x^2(u) du = 2$ and $\int_0^\infty ux'(u)^2 du = \frac{1}{2}$. Substituting this into (4.5), we obtain the bound $a^* \leq \frac{3}{2}\pi^{1/3} = 2.1968\dots$. \square

4.2. *Proof of Proposition 4.* Let P_h be the distribution of standard one-dimensional Brownian motion $(B_t)_{t \geq 0}$ conditioned on starting at h and let E_h be the corresponding expectation. Define

$$(4.9) \quad T_u = \inf\{t \geq 0: B_t = u\}, \quad u \in \mathbb{R}.$$

Note that the following lemma in particular implies Proposition 4(ii).

LEMMA 7. *For every $a \in \mathbb{R}$ and $h, t > 0$,*

$$(4.10) \quad \begin{aligned} w_a(h, t) &= \exp(at) E_{h/2} \left(\exp \left(- \int_0^t 2B_s ds \right) \middle| T_0 = t \right) \varphi_h(t), \\ \varphi_h(t) &= \frac{P_{h/2}(T_0 \in dt)}{dt} = \frac{h}{2\sqrt{2\pi t^3}} \exp \left(-\frac{h^2}{8t} \right). \end{aligned}$$

Consequently,

$$(4.11) \quad y_a(h) = E_{h/2} \left(\exp \left(\int_0^{T_0} (a - 2B_s) ds \right) \right).$$

PROOF. Recall (1.9). According to Ethier and Kurtz (1986), Theorem 6.1.4, the process $(Y_t^*)_{t \geq 0} = (X_{A^{*-1}(t)}^*)_{t \geq 0}$ is a diffusion with generator [see (1.4)]

$$(4.12) \quad (\tilde{G}^* f)(u) = \frac{1}{u} (G^* f)(u) = 2f''(u), \quad f \in C_c^2(\mathbb{R}^+).$$

In other words, the distribution of $(Y_t^*)_{t \geq 0}$ under \mathbb{P}_h^* is equal to that of $(B_{4t \wedge T_0})_{t \geq 0}$ under P_h , which in turn is equal to that of $(2B_{t \wedge T_0})_{t \geq 0}$ under $P_{h/2}$. Thus, noting that $(d/dt)A^{*-1}(t) = 1/X_{A^{*-1}(t)}^*$ and hence $\int_0^{A^{*-1}(t)} X_v^{*2} dv = \int_0^t X_{A^{*-1}(s)}^* ds$, we have

$$(4.13) \quad \begin{aligned} & \mathbb{E}_h^* \left(\exp \left(- \int_0^\infty X_v^{*2} dv \right) \middle| A^*(\infty) = t \right) \\ &= \mathbb{E}_h^* \left(\exp \left(- \int_0^{\xi_0} X_v^{*2} dv \right) \middle| A^*(\xi_0) = t \right) \\ &= \mathbb{E}_h^* \left(\exp \left(- \int_0^{A^{*-1}(t)} X_v^{*2} dv \right) \middle| A^{*-1}(t) = \xi_0 \right) \\ &= E_{h/2} \left(\exp \left(- \int_0^t 2B_s ds \right) \middle| T_0 = t \right), \end{aligned}$$

which proves the first formula in (4.10) [see (2.18)]. In the same way, we see that φ_h defined in (1.10) equals the Lebesgue density of T_0 under $P_{h/2}$, and its explicit shape is given in RY, page 102. Finally, the representation (4.11) is a direct consequence of (2.31). \square

PROOF OF PROPOSITION 4(i). In view of Lemmas 5 and 6, the following lemma implies Proposition 4(i).

LEMMA 8. $z_\alpha = y_\alpha$ for all $\alpha < -2^{1/3}u_1$.

PROOF. Since $y_\alpha(0) = z_\alpha(0) = 1$ and since z_α is bounded on \mathbb{R}_0^+ , it suffices to show that y_α satisfies the same differential equation as z_α [see (4.3)]. But this easily follows from the argument in the proof of KS, Theorem 4.6.4.3, picking (in the notation used there) $\alpha = \alpha < -2^{1/3}u_1$, $k(u) = u$, $\gamma_l = 0$, $b = 0$ and $c = \infty$. \square

PROOF OF PROPOSITION 4(iii) AND (iv). Fix $p \in (1, 2)$ and $q \in (2, \infty)$. Recall (2.32). In the following, c denotes a generic positive constant, possibly varying from line to line.

STEP 1. $W_p^{(1)}$ is integrable at zero.

PROOF. Use (4.10) to estimate $w_{\alpha^*}(h, t) \leq ct^{-3/2}he^{-h^2/8t}$ for any $h \geq 0$ and $t \in (0, 1]$. Use the boundedness of $x_{\alpha^*}^{2-p}$ on \mathbb{R}_0^+ to get

$$\begin{aligned} W_p^{(1)}(t) &\leq c \left(\int_0^\infty h^{1-p} h^p t^{-3p/2} \exp\left(-\frac{ph^2}{8t}\right) dh \right)^{1/p} \\ (4.14) \quad &= ct^{-3/2} \left(\int_0^\infty h \exp\left(-\frac{ph^2}{8t}\right) dh \right)^{1/p} \\ &= ct^{(1/p)-3/2}, \end{aligned}$$

which is integrable at zero. \square

STEP 2. $W_q^{(2)}$ is integrable at zero.

PROOF. As in Step 1, use (4.10) to estimate $w_{\alpha^*}(h, t) \leq ct^{-3/2}he^{-h^2/8t}$, and furthermore use $h^{1+q}e^{-qh^2/16t} \leq ct^{(1+q)/2}$ for any $h \geq 0$ and $t \in (0, 1]$. This gives

$$\begin{aligned} W_q^{(2)}(t) &\leq c \left(\int_0^\infty hx_{\alpha^*}(h)^{2-q} h^q t^{-3q/2} \exp\left(-\frac{qh^2}{8t}\right) dh \right)^{1/q} \\ (4.15) \quad &\leq ct^{-3/2} \left(\int_0^\infty x_{\alpha^*}(h)^{2-q} t^{(1+q)/2} \exp\left(-\frac{qh^2}{16t}\right) dh \right)^{1/q} \\ &= ct^{(1/2q)-1} \left(\int_0^\infty x_{\alpha^*}(h)^{2-q} \exp\left(-\frac{qh^2}{16}\right) dh \right)^{1/q}. \end{aligned}$$

The integral is finite since $\lim_{h \rightarrow \infty} h^{-3/2} \log x_{\alpha^*}(h)$ is finite (see the beginning of Section 2.1). Thus, the r.h.s. of (4.15) is integrable at zero. \square

STEP 3. $W_p^{(1)}$ is integrable at infinity.

PROOF. Since $t \mapsto \frac{1}{2}t^{-3/2}$ is a probability density on $[1, \infty)$, Jensen's inequality and the boundedness of $x_{a^*}^{2-p}$ on \mathbb{R}_0^+ give

$$(4.16) \quad \begin{aligned} \int_1^\infty W_p^{(1)}(t) dt &\leq c \int_1^\infty \frac{1}{2}t^{-3/2} dt \left(\int_0^\infty h^{1-p} t^{3p/2} w_{a^*}(h, t)^p dh \right)^{1/p} \\ &\leq c \left(\int_1^\infty dt \int_0^\infty dh h^{1-p} t^{(3/2)(p-1)} w_{a^*}(h, t)^p \right)^{1/p}. \end{aligned}$$

Use (4.10), Jensen's inequality and the Brownian scaling property to estimate

$$(4.17) \quad \begin{aligned} w_{a^*}(h, t)^p &\leq \varphi_h(t)^{p-1} \varphi_h(t) \mathbf{E}_{h/2} \left(\exp \left(a^* pt - p \int_0^t 2B_s ds \right) \middle| T_0 = t \right) \\ &\leq ch^{p-1} t^{-(3/2)(p-1)} \varphi_{hp^{1/3}}(tp^{2/3}) \\ &\quad \times \mathbf{E}_{(hp^{1/3})/2} \left(\exp \left(a^* p^{1/3} tp^{2/3} - \int_0^{tp^{2/3}} 2B_s ds \right) \middle| T_0 = tp^{2/3} \right) \\ &= ch^{p-1} t^{-(3/2)(p-1)} w_{a^* p^{1/3}}(hp^{1/3}, tp^{2/3}). \end{aligned}$$

Substitute this into (4.16), recall (2.31) and use Lemmas 6 and 8, to get

$$(4.18) \quad \left(\int_1^\infty W_p^{(1)}(t) dt \right)^p \leq c \int_0^\infty z_{a^* p^{1/3}}(hp^{1/3}) dh.$$

The r.h.s. is finite by (4.4). \square

STEP 4. $W_q^{(2)}$ is integrable at infinity if $q \in (2, \infty)$ is sufficiently close to 2.

PROOF. Estimate in the same way as in (4.16) and (4.17), but do not estimate $x_{a^*}^{2-q}(h)$. The result is

$$(4.19) \quad \left(\int_4^\infty W_q^{(2)}(t) dt \right)^q \leq c \int_0^\infty h^q x_{a^*}(h)^{2-q} z_{a^* q^{1/3}}(hq^{1/3}) dh.$$

For q sufficiently close to 2 we have $a^* q^{1/3} < -2^{1/3} u_1$ (see Lemma 6), and so we may apply (4.4). Combine the latter with the fact that $\lim_{h \rightarrow \infty} h^{-3/2} \log x_{a^*}(h)$ is finite to deduce that the r.h.s. of (4.19) is finite for q sufficiently close to 2. \square

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